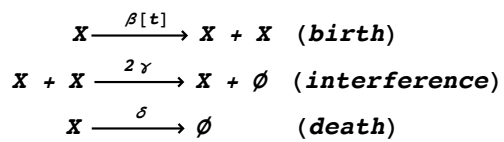


Interference competition and noise

Model 1

Individual model :



modelled as independent Poisson processes and assuming mass - action when appropriate.

Environmental noise :

$$\beta[t] := \beta_0 e^{\theta[t]}$$

$$d\theta = -a \theta dt + b dW$$

with $a > 0$.

Population model :

$$\begin{cases} dX = (\beta_0 e^\theta X - \delta X - \gamma X^2) dt \\ d\theta = -a \theta dt + b dW \end{cases}$$

with $a > 0$.

Linear approximation :

$$\bar{X} = (\beta_0 - \delta) / \gamma$$

$$U = X - \bar{X}$$

$$\begin{cases} dU = -(\beta_0 - \delta) U dt + \beta_0 (\beta_0 - \delta) \gamma^{-1} \theta dt \\ d\theta = -a \theta dt + b dW \end{cases}$$

with $\beta_0 > \delta$ and $a > 0$.

Ornstein - Ulenbeck process :

$U \sim N[0, \sigma^2]$ and hence $X \sim N[\bar{X}, \sigma^2]$ with $\bar{X} = (\beta_0 - \delta) \gamma^{-1}$

$$S_1[\omega] = \frac{(\beta_0 - \delta)^2 \beta_0^2}{\gamma^2 (\omega^2 + (\beta_0 - \delta)^2)} \frac{b^2}{\omega^2 + a^2}$$

$$C_1[t] = \begin{cases} \frac{b^2 \beta_0^2 (\beta_0 - \delta) e^{-t(a+\beta_0-\delta)} (a e^{t a} - (\beta_0 - \delta) e^{t(\beta_0 - \delta)})}{2 a \gamma^2 (a^2 - (\beta_0 - \delta)^2)} & \text{if } \beta_0 \neq a + \delta \\ \frac{b^2 e^{-a t} (1 + a t) (a + \delta)^2}{4 a \gamma^2} & \text{if } \beta_0 = a + \delta \end{cases}$$

for $t > 0$, and so, in particular,

$$\sigma_1^2 = \begin{cases} \frac{b^2 \beta_0^2 (\beta_0 - \delta)}{2 a \gamma^2 (a + \beta_0 - \delta)} & \text{if } \beta_0 \neq a + \delta \\ \frac{b^2 (a + \delta)^2}{4 a \gamma^2} & \text{if } \beta_0 = a + \delta \end{cases}$$

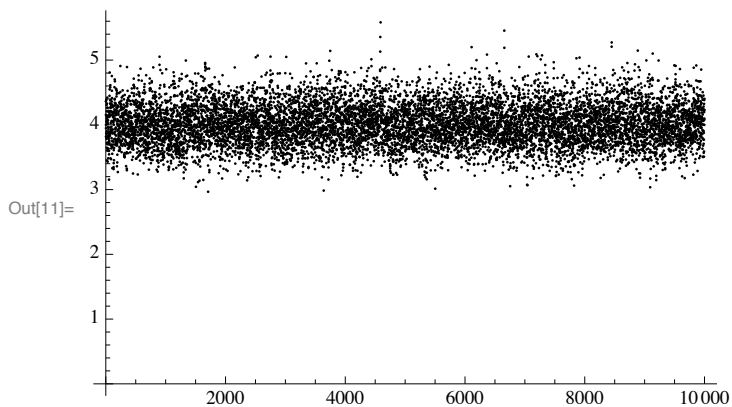
```

In[1]:= << MultivariateStatistics`;

 $\beta_0 = 5$ ;  $\gamma = 1$ ;  $\delta = 1$ ;  $a = 1$ ;  $b = .1$ ;
(*  $\beta_0=2$ ; *)
SeedRandom[1234];
tMax = 10000;
dt = .01;
t = 0;
 $x = (\beta_0 - \delta) / \gamma$ ;
 $\theta = 0$ ;
data1 = {};
While[
  t < tMax,
  Do[
    z = RandomReal[NormalDistribution[0, 1]] // N;
     $d\theta = -a \theta dt + \sqrt{dt} b z$ ;
     $dx = (\beta_0 e^\theta x - \delta x - \gamma x^2) dt$ ;
    t = t + dt;
     $\theta = \theta + d\theta$ ;
     $x = x + dx$ ,
    {i, 1, 100}];
  data1 = Join[data1, {{t, x}}];

In[11]:= ListPlot[data1, PlotStyle -> {Black, PointSize[.001]}, AxesOrigin -> {0, 0}]

```

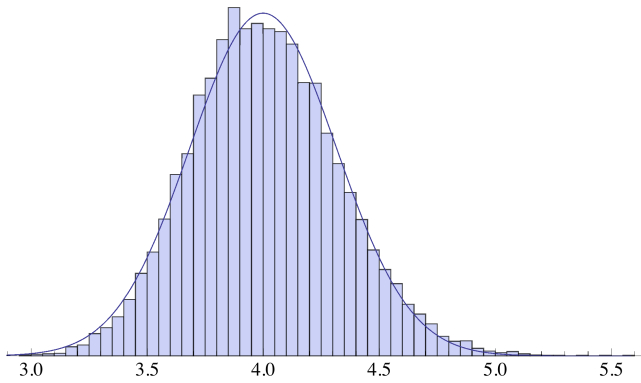


```
In[12]:=  $\mu = (\beta_0 - \delta) / \gamma;$ 
```

$$\sigma = \sqrt{\text{If}[\beta_0 \neq a + \delta, \frac{b^2 \beta_0^2 (\beta_0 - \delta)}{2 a \gamma^2 (a + \beta_0 - \delta)}, \frac{b^2 (a + \delta)^2}{4 a \gamma^2}];}$$

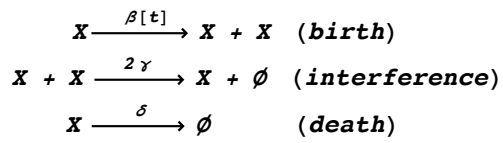
```
distr1 =
Show[
Histogram[data1[[All, 2]], Automatic, "ProbabilityDensity"],
Plot[PDF[NormalDistribution[ $\mu$ ,  $\sigma$ ], x], {x, 0, Max[data1[[All, 2]]]}, PlotRange -> All,
AxesOrigin -> {0, 0}
]
```

```
Out[14]=
```



Model 2

Individual model :



modelled as independent Poisson processes and assuming mass - action when appropriate.

Environmental noise :

$$\beta[t] := \beta_0 + \epsilon \xi[t]$$

where $\xi[t]$ is the standard white noise.

Population model :

$$dX = (\beta_0 X - \delta X - \gamma X^2) dt + \epsilon X dW$$

Linear approximation :

$$\bar{X} = \frac{(\beta_0 - \delta)}{\gamma}$$

$$U = X - \bar{X}$$

$$dU = -(\beta_0 - \delta) U dt + \epsilon \frac{(\beta_0 - \delta)}{\gamma} dW$$

Ornstein - Ulenbeck process :

$U \sim N[0, \sigma^2]$ and hence $X \sim N[\bar{X}, \sigma^2]$ with $\bar{X} = (\beta_0 - \delta) \gamma^{-1}$

$$S_2[\omega] = \frac{\epsilon^2 (\delta - \beta_0)^2}{\gamma^2 (\omega^2 + (\beta_0 - \delta)^2)}$$

$$2[t] = \frac{\epsilon^2 (\beta_0 - \delta)}{2 \gamma^2} e^{-(\beta_0 - \delta) t}$$

for $t > 0$, and in particular

$$\sigma_2^2 = \frac{\epsilon^2 (\beta_0 - \delta)}{2 \gamma^2}$$

Since we have no natural interpretation of the ϵ , we have to calibrate the variances of the two models such that they become equal, which gives

$$\epsilon^2 = \text{If} \left[\beta_0 \neq a + \delta, \frac{b^2 \beta_0^2}{a (a + \beta_0 - \delta)}, \frac{b^2 (a + \delta)^2}{2 a (\beta_0 - \delta)} \right]$$

```

In[15]:= << MultivariateStatistics` ;

β0 = 5; γ = 1; δ = 1; a = 1; b = .1;
(* β0=2; *)

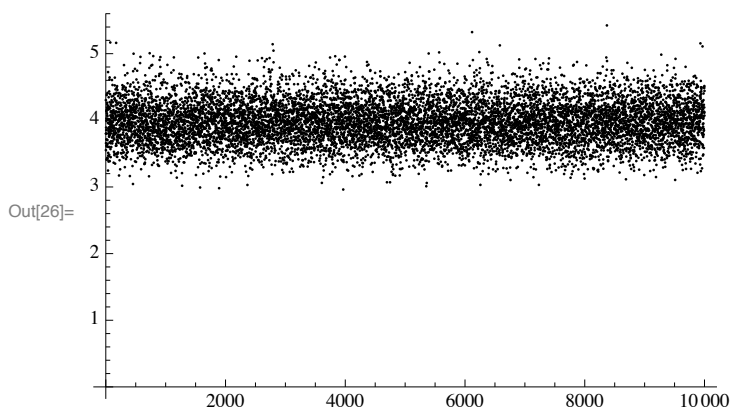
ε =  $\sqrt{\text{If}[\beta_0 \neq a + \delta, \frac{b^2 \beta_0^2}{a(a + \beta_0 - \delta)}, \frac{b^2 (a + \delta)^2}{2a(\beta_0 - \delta)}]}$ ;

(* this ε equalizes the vars of the two models *)

SeedRandom[1234];
tMax = 10 000;
dt = .01;
t = 0;
x = (β0 - δ) / γ;
θ = 0;
data2 = {};
While[
  t < tMax,
  Do[
    z = RandomReal[NormalDistribution[0, 1]] // N;
    dx = (β0 x - δ x - γ x2) dt + ε x  $\sqrt{dt}$  z;
    t = t + dt;
    x = x + dx,
    {i, 1, 100}];
  data2 = Join[data2, {{t, x}}];

In[26]:= ListPlot[data2, PlotStyle → {Black, PointSize[.001]}, AxesOrigin → {0, 0}]

```



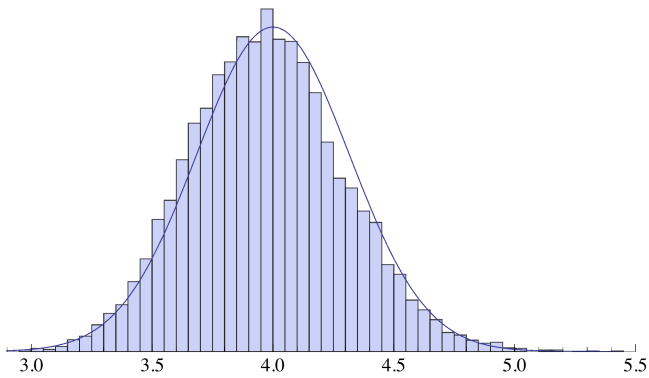
```
In[27]:=  $\mu = (\beta_0 - \delta) / \gamma;$ 
```

$$\sigma = \sqrt{\frac{\epsilon^2 (\beta_0 - \delta)}{2 \gamma^2}};$$

```
distr2 =
```

```
Show[  
  Histogram[data2[[All, 2]], Automatic, "ProbabilityDensity"],  
  Plot[PDF[NormalDistribution[ $\mu$ ,  $\sigma$ ], x], {x, 0, Max[data2[[All, 2]]]}, PlotRange → All],  
  AxesOrigin → {0, 0}  
]
```

```
Out[29]=
```



Comparing models 1 and 2

The difference between the models is not in the stationary distribution, but in the time structure

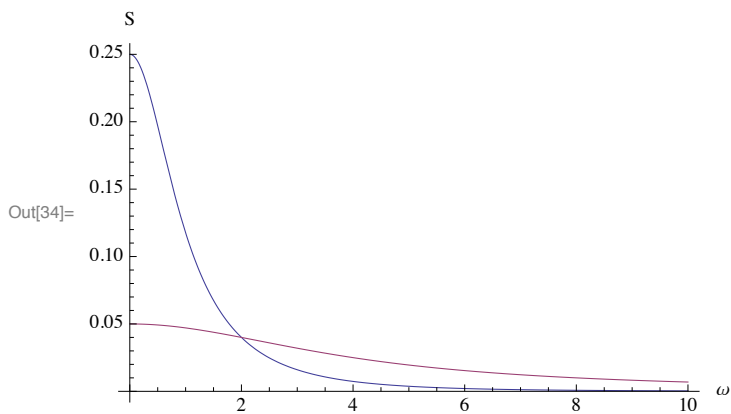
```
In[30]:=  $\beta_0 = 5; \gamma = 1; \delta = 1; a = 1; b = .1;$   
(*  $\beta_0=2;$  *)
```

$$\epsilon = \sqrt{\text{If}[\beta_0 \neq a + \delta, \frac{b^2 \beta_0^2}{a(a + \beta_0 - \delta)}, \frac{b^2 (a + \delta)^2}{2a(\beta_0 - \delta)}]}$$

```
In[32]:=  $S_1[\omega_] := \frac{b^2 (\delta - \beta_0)^2 \beta_0^2}{\gamma^2 (a^2 + \omega^2) (\omega^2 + (\delta - \beta_0)^2)};$ 
```

$$S_2[\omega_] := \frac{\epsilon^2 (\delta - \beta_0)^2}{\gamma^2 (\omega^2 + (\beta_0 - \delta)^2)}$$

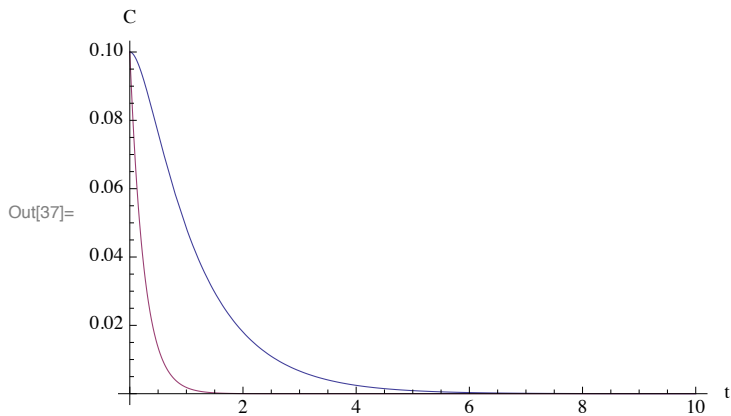
```
Plot[{S1[ $\omega$ ], S2[ $\omega$ ]}, { $\omega$ , 0, 10}, PlotRange -> All, AxesLabel -> {" $\omega$ ", "S"}]
```



```
In[35]:= Cov1[t_] :=
  If[β0 ≠ a + δ,  $\frac{b^2 \beta_0^2 (\beta_0 - \delta) e^{-t(a+\beta_0-\delta)} (a e^{t a} - (\beta_0 - \delta) e^{t(\beta_0-\delta)})}{2 a \gamma^2 (a^2 - (\beta_0 - \delta)^2)}$ ,  $\frac{b^2 e^{-a t} (1 + a t) (a + \delta)^2}{4 a \gamma^2}$ ];
```

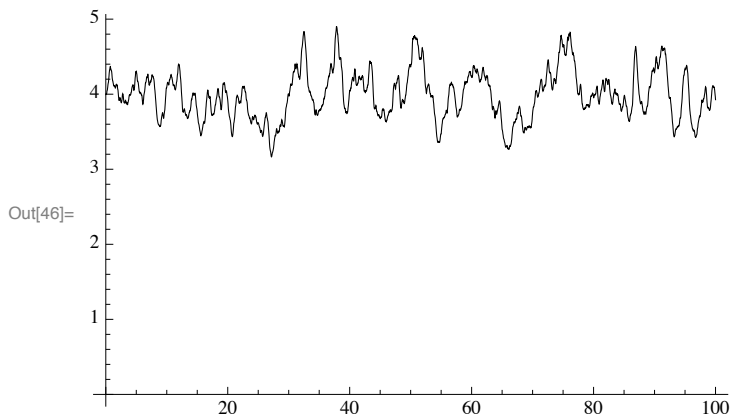
```
Cov2[t_] :=  $\frac{\epsilon^2 (\beta_0 - \delta)}{2 \gamma^2} e^{-(\beta_0 - \delta) t}$ ;
```

```
Plot[{Cov1[t], Cov2[t]}, {t, 0, 10}, PlotRange → All, AxesLabel → {"t", "C"}]
```



```
In[38]:= SeedRandom[1234];
tMax = 100;
dt = .005;
t = 0;
x = (β0 - δ) / γ;
θ = 0;
data1 = {};
While[
  t < tMax,
  Do[
    z = RandomReal[NormalDistribution[0, 1]] // N;
    dθ = -a θ dt + √dt b z;
    dx = (β0 eθ x - δ x - γ x2) dt;
    t = t + dt;
    θ = θ + dθ;
    x = x + dx,
    {i, 1, 1}]; (* collect full data *)
  data1 = Join[data1, {{t, x}}];
```

```
ListPlot[data1, PlotStyle → {Black}, Joined → True, AxesOrigin → {0, 0}]
```

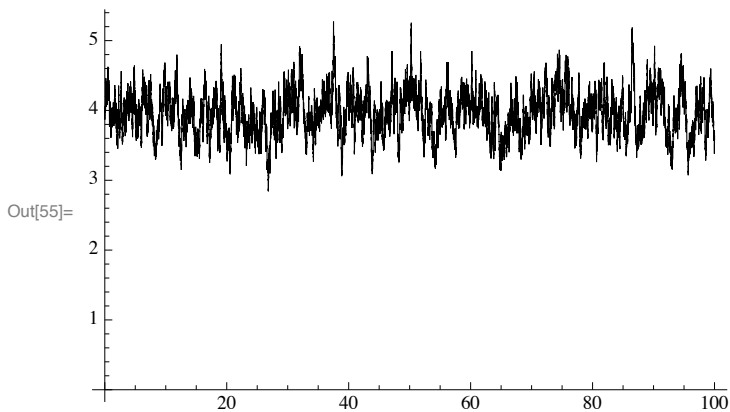


```

In[47]:= SeedRandom[1234];
tMax = 100;
dt = .005;
t = 0;
x = ( $\beta_0 - \delta$ ) /  $\gamma$ ;
 $\theta = 0$ ;
data2 = {};
While[
  t < tMax,
  Do[
    z = RandomReal[NormalDistribution[0, 1]] // N;
    dx = ( $\beta_0 x - \delta x - \gamma x^2$ ) dt +  $\epsilon x \sqrt{dt}$  z;
    t = t + dt;
    x = x + dx,
    {i, 1, 1}]; (* collect full data *)
  data2 = Join[data2, {{t, x}}];

ListPlot[data2, PlotStyle → {Black}, Joined → True, AxesOrigin → {0, 0}]

```



```

In[56]:= Row[{
  Show[
    DensityPlot[PDF[NormalDistribution[ $\mu$ ,  $\sigma$ ], x], {t, 0, tMax}, {x,  $\mu - 4\sigma$ ,  $\mu + 4\sigma$ },
    ListPlot[data1, PlotStyle → {Black}, Joined → True], ImageSize → Small],
  Show[
    DensityPlot[PDF[NormalDistribution[ $\mu$ ,  $\sigma$ ], x], {t, 0, tMax}, {x,  $\mu - 4\sigma$ ,  $\mu + 4\sigma$ },
    ListPlot[data2, PlotStyle → {Black}, Joined → True], ImageSize → Small]
}]

```

