

# ON THE PROBABILITY OF INVASION IN A MULTI-TYPE BRANCHING PROCESS WITH A SINGLE BIRTH STATE

STEFAN A. H. GERITZ

Consider a multi-type branching process with states  $0, \dots, n$ , and where 0 corresponds to the unique birth state, and let  $b_j$  denote the birth rate and  $d_j$  the death rate in state  $j$ , and let  $t_{ij}$  be the transition rate from state  $j$  to state  $i$ . For the conservation of probability mass we necessarily have

$$(1) \quad t_{jj} = - \sum_{i \neq j} t_{ij} \quad \forall j.$$

Let further  $p_j[l]$  denote the probability that an individual presently in state  $j$  will produce  $l$  offspring during the rest of its stay in the same state  $j$ , and let  $q_j(k)$  denote the probability that an individual presently in state  $j$  will produce  $k$  offspring during the rest of its life in the present state and all other states it will visit thereafter. Then

$$(2) \quad p_j[l] = \left( \frac{b_j}{b_j + d_j - t_{jj}} \right)^l \left( \frac{d_j - t_{jj}}{b_j + d_j - t_{jj}} \right)$$

(i.e., the probability that there are  $l$  birth-events followed by a single non-birth event which terminates the stay in state  $j$  either by a death event or a transition to another state), and

$$(3) \quad q_j[k] = p_j[k] \frac{d_j}{d_j - t_{jj}} + \sum_{l=0}^k p_j[l] \sum_{i \neq j} \left( q_i[k-l] \frac{t_{ij}}{d_j - t_{jj}} \right)$$

(i.e., the probability of producing  $k$  offspring in state  $j$  followed by a death event plus the probability of producing  $l$  offspring in state  $j$  and  $k-l$  offspring during the rest of the individual's life after a transition to another state).

Let  $f_j(z)$  and  $g_j(z)$  denote the probability generating functions of the distributions  $\{p_j[l]\}_{l \geq 0}$  and  $\{q_j[k]\}_{k \geq 0}$ . Then

$$(4) \quad f_j(z) = \frac{d_j - t_{jj}}{(1-z)b_j + d_j - t_{jj}}$$

and after some pretty straightforward calculations, also involving equation (4),

$$(5) \quad g_j(z) \left( (1-z)b_j + d_j \right) = d_j + \sum_{\forall i} g_i(z) t_{ij}.$$

---

*Date:* 25 May 2009.

Differentiation of equation (5) gives

$$(6) \quad R_j d_j - b_j = \sum_{\forall i} R_i t_{ij}$$

where we used that  $g_j(1) = 1$  and  $g'_j(1) = \mathcal{E}_j\{k\} = R_j$ , which is the reproduction ratio of state  $j$ . Note that in particular  $R_0$  is the well-known *basic reproduction ratio*. Define

$$(7) \quad \mathbf{R} := \begin{pmatrix} R_0 & \dots & R_n \\ 0 & \dots & 0 \\ \vdots & & \vdots \\ 0 & \dots & 0 \end{pmatrix}, \quad \mathbf{B} := \begin{pmatrix} b_0 & \dots & b_n \\ 0 & \dots & 0 \\ \vdots & & \vdots \\ 0 & \dots & 0 \end{pmatrix}$$

$$\mathbf{D} := \begin{pmatrix} d_0 & \dots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \dots & d_n \end{pmatrix}, \quad \mathbf{T} := \begin{pmatrix} t_{00} & \dots & t_{0n} \\ \vdots & & \vdots \\ t_{n0} & \dots & t_{nn} \end{pmatrix}$$

Since there is only one birth state,  $\mathbf{R}$  is equal to the so-called next generation matrix. Equation (6) can be written in matrix notation as

$$(8) \quad \mathbf{R}(\mathbf{D} - \mathbf{T}) = \mathbf{B}$$

or equivalently

$$(9) \quad \mathbf{R} = \mathbf{B}(\mathbf{D} - \mathbf{T})^{-1}$$

which is possible because  $\mathbf{D} - \mathbf{T}$  is strictly diagonally dominant and thus can be inverted.

Next, let  $z_j$  denote the probability of the eventual extinction of the branching process starting in state  $j$ . Then, substitution of  $z = z_0$  in equation (5) gives

$$(10) \quad z_j((1 - z_0)b_j + d_j) = d_j + \sum_{\forall i} z_i t_{ij}$$

where we used that  $g_j(z_0) = z_j$  for all  $j$ . Let  $\pi_j = 1 - z_j$  denote the probability of invasion starting from state  $j$ , then from equation (10) and equation (1) we get that

$$(11) \quad \pi_j(\pi_0 b_j + d_j) = \pi_0 b_j + \sum_{\forall i} \pi_i t_{ij}.$$

Define

$$(12) \quad \mathbf{\Pi} := \begin{pmatrix} \pi_0 & \dots & \pi_n \\ 0 & \dots & 0 \\ \vdots & & \vdots \\ 0 & \dots & 0 \end{pmatrix}$$

then equation (11) can be rewritten as

$$(13) \quad \mathbf{\Pi}(\pi_0 \mathbf{B} + \mathbf{D} - \mathbf{T}) = \pi_0 \mathbf{B}.$$

Right-multiplication with  $(\mathbf{D} - \mathbf{T})^{-1}$ , using equation (9), subsequently gives

$$(14) \quad \mathbf{\Pi}(\pi_0 \mathbf{R} + \mathbf{I}) = \pi_0 \mathbf{R}$$

or equivalently,

$$(15) \quad \mathbf{\Pi} = \pi_0 \mathbf{R}(\pi_0 \mathbf{R} + \mathbf{I})^{-1}$$

where  $\mathbf{I}$  is the identity matrix. We can do this because  $\pi_0 \mathbf{R} + \mathbf{I}$  is the product of two non-singular matrices, namely  $\pi_0 \mathbf{B} + \mathbf{D} - \mathbf{T}$ , which is strictly diagonally dominant, and  $(\mathbf{D} - \mathbf{T})^{-1}$ . Hence  $\pi_0 \mathbf{R} + \mathbf{I}$  is non-singular itself and can be inverted. Formal expansion of the right hand side of equation (15) gives

$$(16) \quad \mathbf{\Pi} = \pi_0 \mathbf{R} \sum_{i=0}^{\infty} (-1)^i \pi_0^i \mathbf{R}^i$$

which converges whenever all eigenvalues of  $\pi_0 \mathbf{R}$  lie inside the unit circle in the complex plane, i.e., whenever  $\pi_0 R_0 < 1$ . Writing out equation (18) for the upper leftmost element (i.e., the only element that matters, really), we get

$$(17) \quad \pi_0 = \frac{\pi_0 R_0}{1 + \pi_0 R_0}$$

i.e.,  $\pi_0 = 0$  or

$$(18) \quad \pi_0 = \frac{R_0 - 1}{R_0}$$

whenever the latter is positive, i.e., whenever  $R_0 > 1$ . If  $R_0 \leq 0$ , then  $\pi_0 = 0$  is the only solution. Thus, in conclusion, we have shown that

$$(19) \quad \pi_0 = \begin{cases} 0 & \text{if } R_0 \leq 0 \\ \frac{R_0 - 1}{R_0} & \text{if } R_0 > 1. \end{cases}$$

I would like to emphasize that this result is exact.