

HU, Probability Theory Fall 2015, Problems 7 (28.10.2015)

In the problems all random variables live in a probability space (Ω, \mathcal{F}, P) .

1. On a probability space (Ω, \mathcal{F}, P) ,

let $(X_n(\omega) : n \in \mathbb{N})$ be a sequence of exponential random variables such that

$$P(X_1 > t_1, \dots, X_n > t_n) = \exp\left(-\lambda \sum_{i=1}^n t_i\right) \quad \forall n \in \mathbb{N}, t_1, \dots, t_n \geq 0,$$

where $\lambda > 0$ is a parameter.

(a) Show that the random variables $(X_n(\omega) : n \in \mathbb{N})$ are independent under P .

Solution

$$P(X_1 > t_1, \dots, X_n > t_n) = \exp\left(-\lambda \sum_{i=1}^n t_i\right) = \prod_{i=1}^n \exp\left(-\lambda t_i\right) = \prod_{i=1}^n P(X_i > t_i)$$
$$\forall n \in \mathbb{N}, t_1, \dots, t_n \geq 0,$$

and since the collection of rectangles

$$\mathcal{I} = \left\{ (t_1, 1] \times (t_2, 1] \times \dots \times (t_n, 1] : t_i \in [0, 1] \right\}$$

form a π -system (closed under intersections) which generate the borel σ -algebra $\mathcal{B}([0, 1]^n) = \mathcal{B}([0, 1])^{\otimes n}$, by Dynkin lemma on uniqueness of probability measures,

$$P(X_1 \in B_1, \dots, X_n \in B_n) = \prod_{i=1}^n P(X_i \in B_i) \quad \forall B_i \in \mathcal{B}([0, 1]), 1 \leq i \leq n$$

which means that X_1, \dots, X_n are P -independent, and they are also identically distributed with λ -exponential distribution.

(b) Let

$$Y_n(\omega) := \min\{X_1(\omega), X_2(\omega), \dots, X_n(\omega)\}.$$

Compute $P(Y_n > t)$, and compute also the probability density function of Y_n .

Solution

$$P(Y_n > t) = P(X_1 > t, \dots, X_n > t) = \prod_{i=1}^n P(X_i > t) = \exp(-\lambda t)^n = \exp(-n\lambda t)$$

and the probability density function of Y_n is given by

$$f_{Y_n}(t) := -\frac{d}{dt}P(Y_n > t) = -\frac{d}{dt}\exp(-n\lambda t) = n\lambda \exp(-n\lambda t)$$

which is the density of a λ -exponential random variable with parameter $n\lambda$.

(c) Let $X_n^*(\omega) = \max\{X_1(\omega), X_2(\omega), \dots, X_n(\omega)\}$

Compute $P(X_n^* \leq t)$. Compute also the probability density function of X_n^* . **Solution**

$$P(X_n^* \leq t) = P(X_1 \leq t, X_2 \leq t, X_n \leq t) = \prod_{i=1}^n P(X_i \leq t) = (1 - \exp(-\lambda t))^n$$

and the probability density function of X_n^* is given by

$$f_{X_n^*}(t) := \frac{d}{dt}P(X_n^* \leq t) = n\lambda(1 - \exp(-\lambda t))^{n-1} \exp(-\lambda t)$$

(d) Compute $\lim_{n \rightarrow \infty} P(\lambda X_n^* \leq t + \log(n))$.

Hint: $(1 + x/n)^n \rightarrow \exp(x)$ as $n \rightarrow \infty$.

Solution

$$P(\lambda X_n^* \leq t + \log(n)) = (1 - \exp(-t - \log(n)))^n = \left(1 - \frac{\exp(-t)}{n}\right)^n \tag{0.1}$$

$$\rightarrow \exp(-\exp(-t))$$

as $n \rightarrow \infty$. The function

$$G(t) = \exp(-\exp(-t))$$

is continuous and strictly increasing, with $G(-\infty) = 0$ and $G(+\infty) = 1$, it is the cumulative distribution function of a probability distribution on \mathbb{R} . Such distribution is called Gumbel extreme value distribution, which in the context of extreme value theory.

Let $\xi_n(\omega) = (\lambda X^{*n}(\omega) - \log n)$. We have shown that $G_n(t) := P(\xi_n \leq t) \rightarrow G(t)$. When a sequence of cumulative distribution functions $G_n(t)$ converges to a cumulative distribution function $G(t)$ at all points t where $G(t)$ is continuous, we say that ξ_n converges in law to the limiting distribution G .

2. Consider a sequence of random variables $(U_k(\omega) : k \in \mathbb{N})$ such that for $\forall t_1, \dots, t_n \in [0, 1]$,

$$P(U_1 \leq t_1, \dots, U_n \leq t_n) = \prod_{k=1}^n t_k$$

- (a) Show that $(U_k(\omega) : k \in \mathbb{N})$ are independent and uniformly distributed on $[0, 1]$.

Solution

$$P(U_1 \leq t_1, \dots, U_n \leq t_n) = \prod_{k=1}^n t_k = \prod_{k=1}^n P(U_k \leq t_k)$$

which implies as in the previous exercise that U_1, \dots, U_n are P -independent, and also that they are identically distributed where the common probability distribution is Lebesgue measure on the interval $[0, 1]$.

- (b) Consider $\bar{U}_n(\omega) = \max\{U_1(\omega), \dots, U_n(\omega)\}$.

Compute the cumulative distribution function of \bar{U}_n , $F_{\bar{U}_n}(t) = P(\bar{U}_n \leq t)$.

Solution

$$P(\bar{U}_n \leq t) = P(U_1 \leq t, \dots, U_n \leq t) = \prod_{i=1}^n P(U_i \leq t) = t^n$$

- (c) Show that $\lim_{n \rightarrow \infty} \bar{U}_n(\omega) = 1$ \mathbb{P} -almost surely. **Solution** Note that $\forall \omega \in \Omega$,

$$0 \leq \bar{U}_n(\omega) \leq \bar{U}_{n+1}(\omega) \uparrow \bar{U}_\infty(\omega) \leq 1$$

where the limit exists since the sequence $\bar{U}_n(\omega)$ is non-decreasing and bounded by 1 for each ω .

We show that $P(\bar{U}_\infty = 1) = 1$.

Equivalently we will show that

$$0 \stackrel{?}{=} P(\bar{U}_\infty < 1) = P\left(\bigcup_{m \in \mathbb{N}} \{\bar{U}_\infty < 1 - 1/m\}\right)$$

and for each fixed $m \in \mathbb{N}$

$$\{\bar{U}_\infty < 1 - 1/m\} = \bigcap_{n \in \mathbb{N}} \{U_n < 1 - 1/m\}$$

But by the σ -additivity of P ,

$$P\left(\bigcap_{n \in \mathbb{N}} \{U_n < 1 - 1/m\}\right) = \lim_{N \rightarrow \infty} P\left(\bigcap_{1 \leq n \leq N} \{U_n < 1 - 1/m\}\right) = \lim_{N \rightarrow \infty} (1 - 1/m)^N = 0$$

and the claim follows since the countable union of P -null events is a P -null event.

(d) Let $\underline{U}_n(\omega) = \min\{U_1(\omega), \dots, U_n(\omega)\}$.

Compute the cumulative distribution function of \underline{U}_n , $F_{\underline{U}_n}(t) = P(\underline{U}_n \leq t)$.

Solution

$$\begin{aligned} P(\underline{U}_n \leq t) &= 1 - P(\underline{U}_n > t) = 1 - P(U_1 > t, \dots, U_n > t) = \\ &= 1 - \prod_{i=1}^n P(U_i > t) = 1 - (1 - t)^n \end{aligned}$$

(e) Show that $\lim_{n \rightarrow \infty} \underline{U}_n(\omega) = 0$ \mathbb{P} -almost surely.

Hint: $V_n = (1 - U_n)$ has the same distribution as U_n , which implies that \underline{U}_n and $(1 - \bar{U}_n)$ have the same distribution.

Solution

Note that $\forall \omega \in \Omega$,

$$1 \geq \underline{U}_n(\omega) \geq \underline{U}_{n+1}(\omega) \downarrow \underline{U}_\infty(\omega) \geq 0$$

where the limit exists since the sequence $\bar{U}_n(\omega)$ is non-increasing and non-negative for each ω .

We show that $P(\underline{U}_\infty = 0) = 1$.

Equivalently we will show that

$$0 \stackrel{?}{=} P(\underline{U}_\infty > 0) = P\left(\bigcup_{m \in \mathbb{N}} \{\underline{U}_\infty > 1/m\}\right)$$

and for each fixed $m \in \mathbb{N}$

$$\{\underline{U}_\infty > 1/m\} = \bigcap_{n \in \mathbb{N}} \{U_n > 1/m\}$$

But by the σ -additivity of P ,

$$P\left(\bigcap_{n \in \mathbb{N}} \{U_n > 1/m\}\right) = \lim_{N \rightarrow \infty} P\left(\bigcap_{1 \leq n \leq N} \{U_n > 1/m\}\right) = \lim_{N \rightarrow \infty} (1 - 1/m)^N = 0$$

and the claim follows since the countable union of P -null events is a P -null event.

3. (a) let $X(\omega), X_n(\omega), n \in \mathbb{N}$ such that $X_n(\omega) \rightarrow X(\omega)$ P -almost surely. Show that also the Cesaro mean converges P -almost surely to X

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i(\omega) = X(\omega) \quad P\text{-almost surely}$$

Solutions Byt assumption, there exists an N with $P(N) = 0$ such that $\forall \omega \in N^c, \forall \varepsilon > 0 \exists K(\varepsilon, \omega)$ such that $\forall n \geq K(\varepsilon, \omega)$,

$$|X_n(\omega) - X(\omega)| < \varepsilon$$

Let

$$\bar{S}_n(\omega) = \frac{1}{n} \sum_{i=1}^n X_i(\omega)$$

and

$$\bar{S}_n(\omega) - X(\omega) = \frac{1}{n} \sum_{i=1}^n (X_i(\omega) - X(\omega)).$$

By the triangle inequality, when $n \geq K(\varepsilon, \omega)$

$$\begin{aligned} |\bar{S}_n(\omega) - X(\omega)| &\leq \frac{1}{n} \sum_{i=1}^{K(\varepsilon, \omega)-1} |X_i(\omega) - X(\omega)| + \frac{1}{n} \sum_{i=K(\varepsilon, \omega)}^n |X_i(\omega) - X(\omega)| \\ &\leq \frac{1}{n} C(\omega, \varepsilon) + \varepsilon \end{aligned} \tag{0.2}$$

where $C(\omega, \varepsilon)$ does not depend from n . Therefore when $\omega \in N^c$, $\forall \varepsilon > 0$

$$\lim_{n \rightarrow \infty} |\bar{S}_n(\omega) - X(\omega)| \leq \varepsilon$$

and since $\varepsilon > 0$ is arbitrary this means that

$$\lim_{n \rightarrow \infty} \bar{S}_n(\omega) = X(\omega) \quad \forall \omega \in N^c.$$

- (b) Assume now that $E_P(|X_n - X|) \rightarrow 0$, as $n \rightarrow \infty$ (without assuming P -almost sure convergence). We also need to assume that $E_P(|X_n|) < \infty \forall n$, in order to guarantee that the Cesaro means \bar{S}_n are integrable.

Show that the Cesaro mean is converging in $L^1(P)$, that is

$$\lim_{n \rightarrow \infty} E_P \left(\left| \left\{ \frac{1}{n} \sum_{i=1}^n X_i \right\} - X(\omega) \right| \right) \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

Hint: note that by the triangle inequality

$$\begin{aligned} \left| \left\{ \frac{1}{n} \sum_{i=1}^n X_i \right\} - X(\omega) \right| &\leq \frac{1}{n} \sum_{i=1}^n |X_i - X(\omega)| = \\ &\frac{1}{n} \sum_{i=1}^M |X_i - X(\omega)| + \frac{1}{n} \sum_{j=M+1}^n |X_j - X(\omega)| \end{aligned}$$

$\forall n \geq M$, where the inequalities are preserved after taking the expectation.

Solution By assumption $\forall \varepsilon > 0 \exists K(\varepsilon)$ (which now does not depend on ω) such that $\forall n \geq K(\varepsilon)$,

$$E(|X_n - X|) < \varepsilon$$

By the triangle inequality

$$|\bar{S}_n(\omega) - X(\omega)| \leq \frac{1}{n} \sum_{i=1}^{K(\varepsilon)-1} |X_i(\omega) - X(\omega)| + \frac{1}{n} \sum_{i=K(\varepsilon)}^n |X_i(\omega) - X(\omega)| \leq \frac{1}{n} C(\omega, \varepsilon) + \varepsilon \quad (0.3)$$

where

$$C(\omega, \varepsilon) = \sum_{i=1}^{K(\varepsilon)-1} |X_i(\omega) - X(\omega)|$$

In order to make sure that $C(\omega, \varepsilon)$ is integrable, we assume that $E_P(|X_n|) < \infty \forall n$. This means that $|X| \leq |X_n| + |X - X_n|$ is integrable since $E(|X|) \leq E(|X_n|) + E(|X - X_n|) < \infty$ for n large enough, and since X and X_n are integrable $\forall n$ also $|X - X_n| \leq |X| + |X_n|$ is integrable $\forall n$. Therefore $C(\omega, \varepsilon)$ in (0.3) is integrable and does not depend on n . By taking expectation in (0.3)

$$E(|\bar{S}_n - X|) \leq \frac{1}{n}E(C(\varepsilon)) + \varepsilon$$

with $E(C(\varepsilon)) < \infty$, and

$$\lim_{n \rightarrow \infty} E(|\bar{S}_n - X|) \leq \varepsilon, \quad \forall \varepsilon > 0$$

which implies that $\bar{S}_n \xrightarrow{L^1(P)} X$ in $L^1(P)$ -norm.

4. Let $X(\omega), (X_n(\omega) : n \in \mathbb{N})$, random variables on a probability space (Ω, \mathcal{F}, P) .

Show that if $\forall \varepsilon > 0$

$$\sum_{n=0}^{\infty} P(|X_n(\omega) - X(\omega)| > \varepsilon) < \infty$$

it follows $\lim_{n \uparrow \infty} X_n(\omega) = X(\omega)$ P -almost surely.

Hint: show first that

$$\{\omega : X_n(\omega) \not\rightarrow X(\omega)\} = \bigcup_{k \in \mathbb{N}} \{\omega : |X_n(\omega) - X(\omega)| > k^{-1} \text{ infinitely often} \}$$

and recall Borel-Cantelli's lemma.

Solution Let

$$A_n^m = \{\omega : |X_n(\omega) - X(\omega)| > 1/m\}$$

Since by assumption $\forall m \in \mathbb{N}$

$$\sum_{n \rightarrow \infty} P(A_n^m) < \infty$$

by the first Borel Cantelli lemma $\forall m \in \mathbb{N}$

$$P(\limsup_{n \in \mathbb{N}} A_n^m) = 0$$

and since the countable union of P -null events is a P -null event

$$0 = P\left(\bigcup_m \bigcap_k \bigcup_{n \geq k} A_n^m\right)$$

For the complement of this event we have

$$\begin{aligned} 1 &= P\left(\bigcap_{m \in \mathbb{N}} \bigcup_{k \in \mathbb{N}} \bigcap_{n \geq k} (A_n^m)^c\right) \\ &= P\left(\left\{\omega : \forall m \exists K = K(\omega, m) \text{ such that } \forall n \geq K(\omega, m) |X_n(\omega) - X(\omega)| \leq 1/m\right\}\right) \\ &= P\left(\left\{\omega : \lim_{n \rightarrow \infty} X_n(\omega) = X(\omega)\right\}\right) \end{aligned}$$

5. Consider a random variable $X(\omega)$ with $E_P(|X|) < \infty$. Show that

$$E_P(|X| \mathbf{1}(|X| > n)) = \int_{\Omega} |X(\omega)| \mathbf{1}(|X(\omega)| > n) P(d\omega) \rightarrow 0 \text{ as } n \rightarrow \infty.$$

Note that $\forall \omega \in \Omega$,

$$\lim_{n \rightarrow \infty} |X(\omega)| \mathbf{1}(|X(\omega)| > n) = 0$$

simply because $X(\omega) \in \mathbb{R}$, and $\mathbf{1}(|X(\omega)| > n) = 0$ for all $n \geq |X(\omega)|$.

Note also that $\forall \omega \in \Omega$

$$0 \leq |X(\omega)| \mathbf{1}(|X(\omega)| > n) \leq |X(\omega)|$$

where the upper bound is integrable by the assumption $E(|X|) < \infty$. Therefore Lebesgue dominated convergence Theorem applies and we can change the order of limit and expectation obtaining

$$\lim_{n \rightarrow \infty} E_P(|X| \mathbf{1}(|X| > n)) = E_P\left(|X| \lim_{n \rightarrow \infty} \mathbf{1}(|X| > n)\right) = 0$$