## UH, Probability Theory Fall 2015, Solutions Problems 6 (14.10.2015)

1. When the cumulative distribution function  $F_X(t) = P(X \le t)$  of a  $\mathbb{R}$ -valued random variable X which is absolutely continuous with respect to Lebesgue measure, which means

$$F_X(b) = F_X(a) + \int_a^b f_X(t)dt$$

for some Borel measurable function  $f_X(t) \ge 0$ , which is called probability density function. When the classical derivative  $\frac{dF_X}{dt}(t)$  exists at all t, then it is a probability density function. More in general  $\frac{dF_X}{dt}(t) = \frac{dP_X}{dt}(t) =$  is understood as the Radon Nikodym derivative of the push-forward probability measure  $P_X$  with respect to Lebesgue measure.

In such case, for every non-negative and Borel measurable test function  $g(x) \ge 0$  we have

$$E_P(g(X)) = \int_{\Omega} g(X(\omega))P(d\omega) = \int_{\mathbb{R}} g(t)P_X(dt)$$
$$= \int_{\mathbb{R}} g(t)F(dt) = \int_{\mathbb{R}} g(t)f_X(t)dt \qquad (0.1)$$

where  $P_X(B) = \mathbb{P}(\{\omega : X(\omega) \in B\})$  is the pushforward measure of  $\mathbb{P}$  by the random variable X. The integral w.r.t.  $P_X$  on  $\mathbb{R}$  is the same as the Lebesgue Stieltjes integral w.r.t dF, meaning that  $P_X$  coincides with the measure induced by the cumulative distribution function F(t) on  $\mathbb{R}$ .

**Hint:** One possible strategy for this proof is to use the monotone class theorem: Define the class

 $\mathcal{C} = \{g : \mathbb{R} \to [0, \infty) \text{ bounded and Borel measurable such that } (0.3) \text{ holds } \}$ 

and show that C is a monotone class (use the linearity of the integral together with the monotone convergence theorem) which contains the indicators  $\mathbf{1}_{(a,b]}(t) \ \forall a \leq b \in \mathbb{R}$ .

**Solution** The assumption of almost everywhere (w.r.t. Lebesgue measure) differentiability of  $F_X(t)$  is equivalent to

$$F(b) = F(a) + \int_{a}^{b} f_X(t)dt \quad \forall a, b \in \mathbb{R} .$$

$$(0.2)$$

When  $g(t) = \mathbf{1}_{(a,b]}(t)$  we have

$$E_P(g(X)) = P(X \in (a,b]) = F(b) - F(a) = \int_{\mathbb{R}} \mathbf{1}_{(a,b]}(t) f_X(t) dt$$

which means that  $\mathbf{1}_{(a,b]} \in \mathcal{C}$ . If we show that  $\mathcal{C}$  is a monotone class, it follows by the monotone class theorem that  $\mathcal{C}$  contains all functions which are bounded and measurable w.r.t. the  $\sigma$ -algebra generated by the intervals  $\sigma((a,b]: a \leq b \in \mathbb{R}) = \mathcal{B}(\mathbb{R})$  which is the Borel  $\sigma$ -algebra. Lets check: the constant  $1 \in \mathcal{C}$ , since

$$1 = E_P(1) = F(+\infty) - F(-\infty) = \int_{-\infty}^{+\infty} f_X(t) dt$$

Since all integrals in (0.2) are linear w.r.t the integrand function g, it follows that C is a vector space.

Let  $(g_n(t) : n \in \mathbb{N}) \subseteq \mathcal{C}$  with  $0 \leq g_n(t) \uparrow g(t) \leq K < \infty$ . Then

$$E_P(g_n(X)) = \int_{\Omega} g(X(\omega))P(d\omega) = \int_{\mathbb{R}} g_n(t)P_X(dt)$$
$$= \int_{\mathbb{R}} g_n(t)F(dt) = \int_{\mathbb{R}} g_n(t)f_X(t)dt \qquad (0.3)$$

and since the monotone convergence theorem holds for all the integrals in (0.3) it follows that we can take the limit inside the integral and (0.3) holds for g.

When  $g(t) \ge 0$  a non-negative Borel measurable function which is not bounded, let  $g^{(N)}(t) = g(t) \land N$ , then (0.3) holds for each  $g^{(N)}(t)$  and once again the monotone convergence theorem implies that (0.3) holds for g(t) as well.

2. Linearity of the expectation The expectation of a random variable  $X(\omega)$  is defined as

$$E_{\mathbb{P}}(X) = E_{\mathbb{P}}(X^{+}) - E_{\mathbb{P}}(X^{-})$$

where  $X^+ = \max\{X, 0\} \ge 0$ ,  $X^- = \max\{-X, 0\} \ge 0$  are non-negative random variables, and we have defined first for non-negative random variables

$$E_{\mathbb{P}}(X) = \sup_{Y \in \mathcal{S}F: 0 \le Y \le X} \left\{ E_{\mathbb{P}}(Y) \right\}$$

In this way the expectation is well defined unless

$$E_{\mathbb{P}}(X^+) = E_{\mathbb{P}}(X^-) = +\infty.$$

In the lectures we have shown (first for simple random variables and then by the monotone convergence theorem ) that when  $X(\omega) \ge 0$ ,  $Y(\omega) \ge 0$  *P*-almost surely (outside a P-null set), and  $a, b \ge 0$ 

$$E_{\mathbb{P}}(aX + bY) = aE_{\mathbb{P}}(X) + bE_{\mathbb{P}}(Y) \tag{0.4}$$

Show that linearity holds for any random variables X, Y and  $a, b \in \mathbb{R}$  when the expectations on both left and right sides in (0.4) are finite.

**Hint:** write (aX + bY) using the representations  $X = (X^+ - X^-)$ ,  $Y = (Y^+ - Y^-)$ ,  $a = (a^+ - a^-)$ ,  $b = (b^+ - b^-)$ , and integrate the positive parts and negative parts separately.

## Solution

$$aX + bY = (aX + bY)^{+} - (aX + bY)^{-} = (a^{+} - a^{-})(X^{+} - X^{-}) + (b^{+} - b^{-})(Y^{+} - Y^{-}) = (a^{+}X^{+} + a^{-}X^{-} + b^{+}Y^{+} + b^{-}Y^{-}) - (a^{-}X^{+} + a^{-}X^{+} + b^{-}Y^{+} + b^{-}Y^{+})$$

Then

$$\begin{split} E(aX+bY) = & E(a^{+}X^{+}+a^{-}X^{-}+b^{+}Y^{+}+b^{-}Y^{-}) - E(a^{-}X^{+}+a^{-}X^{+}+b^{-}Y^{+}+b^{-}Y^{+}) \\ = & a^{+}E(X^{+}) + a^{-}E(X^{-}) + b^{+}E(Y^{+}) + b^{-}E(Y^{-}) + \\ & - & a^{-}E(X^{+}) + a^{-}E(X^{+}) + b^{-}E(Y^{+}) + b^{-}E(Y^{+}) \\ = & (a^{+}-a^{-})(E(X^{+}) - E(X^{-})) + (b^{+}-b^{-})(E(Y^{+}) - E(Y^{-})) \\ = & aE(X) + bE(Y) \end{split}$$

where we have used linearity for non-negative random variables with non-negative coefficients.

- 3. Let  $U(\omega)$  be uniformly distributed r.v. with values in [0, 1], such that  $\mathbb{P}(\{U \in (a, b]\}) = (b a)$  for  $0 \le a \le b \le 1$ .
  - (a) Show that the powers  $U(\omega)^z$ , with  $z \in \mathbb{Z}$  (the integers) are random variables. Solution The map  $u \mapsto u^n$  is continuous when  $n \in \mathbb{N}\{0, 1, 2...\}$ , and the  $U(\omega)^n$  is a random variable since a continuous function composed with a measurable map is measurable. When z = -n and  $n \ge 1$ , the map  $u \mapsto u^{-n}$  is not continuous at u = 0, and  $0^{-n} = +\infty$ . Nevertheless  $U^{-n}(\omega)$  is a random variable since  $\{\omega : U^{-n}(\omega) \le t\} = \{\omega : U(\omega) \ge t^{-1/n}\} =$  $U^{-1}([t^{-1/n}, +\infty) \in \mathcal{F}$ , since U is a random variable.

(b) Compute the moments  $E_{\mathbb{P}}(U^z) \in [0, +\infty]$  for  $z \in \mathbb{Z}$ . We distinguish 3 cases:  $z \in \mathbb{N}$ , z = -1 and  $z \leq -2$ . For  $n \geq 0$ 

$$E_P(U^n) = \int_0^1 u^n du = \frac{1}{n+1}$$

$$E_P(1/U) = \int_0^1 \frac{1}{u} du = \log(1) - \log(0) = +\infty$$

which implies

$$E_P(U^{-n}) = \infty \quad \forall n \in \mathbb{N}$$

by comparison since  $0 \leq U^{-1} \leq U^{-n}$  when  $n \geq 1$  and  $0 \leq U \leq 1$ , and by the positivity of the expectation  $+\infty = E_P(U^{-1}) \leq E_P(U^{-n})$ .

(c) Compute the exponential moments  $E_{\mathbb{P}}(\exp(tU))$  for  $t \in \mathbb{R}$ .

$$E_{\mathbb{P}}(\exp(tU)) = \int_0^1 \exp(tu) du = \frac{e^t - 1}{t}$$

Note that this is continuous at t = 0,

$$\lim_{t \to 0} E_{\mathbb{P}}(\exp(tU)) = \lim_{t \to 0} \frac{e^t - 1}{t} = \lim_{t \to 0} \frac{e^t}{1} = 1 = E_P(\exp(t_0U))$$

with  $t_0 = 0$ , where we used l'Hospital rule. This is also a consequence of the bounded convegence theorem, for any in  $t \in [-T, T]$  with T > 0, we have  $0 \leq \exp(tU(\omega)) \leq \exp(TU(\omega))$  with  $E_P(\exp(TU)) < \infty$  which implies that the map  $t \mapsto E_P(\exp(tU))$ is continuous.

(d) Compute the trigonometric moments  $E_{\mathbb{P}}(\cos(2\pi tU))$  and  $E_{\mathbb{P}}(\sin(2\pi tU))$  for  $t \in \mathbb{R}$ .

## Solution

$$E_{\mathbb{P}}\left(\sin(2\pi tU)\right) = \int_0^1 \sin(2\pi tu) du = \frac{1 - \cos(2\pi t)}{2\pi t}$$

Note that since  $|\sin(x)| \leq 1$ , it follows by the dominated convergent theorem that the map  $t \mapsto E_{\mathbb{P}}(\sin(2\pi tU))$  is continuous. Note for example that we have continuity at t = 0:

$$\lim_{t \to 0} E_{\mathbb{P}} \left( \sin(2\pi t U) \right) = \lim_{t \to 0} \frac{1 - \cos(2\pi t)}{2\pi t} = \lim_{t \to 0} \frac{2\pi \sin(2\pi t)}{2\pi} = 0 = E_{\mathbb{P}} \left( \sin(2\pi t_0 U) \right)$$
  
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$$E_{\mathbb{P}}(\cos(2\pi tU)) = \int_0^1 \cos(2\pi tu) du = \frac{\sin(2\pi t)}{2\pi t}$$

Note that since  $|\cos(x)| \leq 1$ , it follows by the dominated convergent theorem that the map  $t \mapsto E_{\mathbb{P}}(\cos(2\pi tU))$  is continuous. Note for example that we have continuity at t = 0:

$$\lim_{t \to 0} E_{\mathbb{P}} \big( \cos(2\pi t U) \big) = \lim_{t \to 0} \frac{\sin(2\pi t)}{2\pi t} = \lim_{t \to 0} \frac{2\pi \cos(2\pi t)}{2\pi} = 1 = E_{\mathbb{P}} \big( \cos(2\pi t_0 U) \big)$$
  
with  $t_0 = 0$ .

4. Let  $f:[0,T] \to \mathbb{R}^+$  be a non-negative and bounded measurable function.

We define its upper and lower Riemann-integrals as follows:

 $J^{+}(f) = \inf \{ I(g) : g \ge f, g \text{ takes finitely many values and is piecewise continuous } \}$  $J^{-}(f) = \sup \{ I(g) : g \leq f, g \text{ takes finitely many values and is piecewise continuous } \}$ 

where the integral I(g) of a piecewise continuous function g taking finitely many values is the usual finite sum.

Note that on the real line, a piecewise continuous simple function taking finitely many values is piecewise constant, with representation

$$g(x) = \sum_{k=1}^{n} a_k \mathbf{1}_{E_i}(x)$$
, with  $I(g) = \sum_{k=1}^{n} a_k \operatorname{length}(E_i)$ 

where  $E_i$  are intervals. In the construction of Lebesgue integral, the general definition uses Borel sets instead of intervals.

We say that f is Riemann integrable when  $J^+(f) = J^-(f)$  which defines the Riemann integral J(f) (it is possible that  $J(f) = +\infty$ ).

(a) Show that when f is Riemann integrable the Riemann integral J(f) coincides with Lebegue integral I(f) defined in the lectures. **Hint** We define the Lebesgue integral I(f) of a Borel measurable non-negative function w.r.t. Lebesgue measure as

 $I(f) = \sup \big\{ I(g) : g \le f, \ g \text{ is measurable and takes finitely many values } \big\}$ 

(b) Show that a non-negative continuous function f is Riemann integrable on the compact set [0, T].

**Hint**: a continuous function uniformly continuous on compact sets. Note that you can approximate uniformly on compacts a continuous function by **piecewise continuous** simple functions. **Solution** A simple piecewise continuous function has representation

$$g(x) = \sum_{k=1}^{n} y_k \mathbf{1}_{E_k}(x)$$

where  $E_k$  are intervals.

Since we are integrating with respect to Lebesgue measure, functions which differ on a set of Lebesgue measure zero have the same integral, so we can assume for example that  $E_k = (a_k, b_k]$ . Denote by SC the class of simple piecewise continuous functions. Then for such g Riemann and Lebesgue integrals coincide with the Riemann sum

$$J(g) = I(g) = \sum_{k=1}^{n} y_k \lambda(E_k) = \sum_{k=1}^{n} y_k (b_k - a_k)$$

Let now  $f(x) \ge 0$  a Borel measurable function, and let  $0 \le g'(x) \le f(x) \le g''(x)$  for some  $g', g'' \in SC^+$ . Since the Lebesgue integral is a positive operator,

$$0 \le J(g') = I(g') \le I(f) \le I(g'\prime) = J(g$$
 )

by taking the supremum over  $g' \in SC^+$ , with  $0 \leq g' \leq f$ , and infinum over  $g' \in SC^+$ , with  $f \leq g'$ , it follows that

$$0 \le J^-(f) \le I(f) \le J^+(f)$$

Therefore when f is Riemann integrable, by definition  $J(f) = J^+(f) = J^-(f) = I(f)$ , and by the sandwhiching argument the Lebegue integral and Riemann integral coincide when the latter exists. However the Lebesgue integral is more general, we show in the next example a Borel function  $f \ge 0$  with well define Lebesgue integral such that  $J^-(f) \le J^+(f)$  and the Riemann integral does not exists.

(c) Let  $f(x) = \mathbf{1}_{\mathbb{Q}}(x)$  where  $\mathbb{Q}$  are the rationals.

Show that f is Borel measurable, but is not Riemann integrable on [0, T].

**Hint** : Show that on a compact interval  $[0,T] J^+(f) = T$  and  $J^-(f) = 0$ .

**Solution** Since the rationals are dense in  $\mathbb{R}$ , it follows that the smallest piecewise continuous simple function which is an upper bound for  $\mathbf{1}_Q(x)$  is the constant function with value 1, and t the biggest piecewise continuous simple function which is a lower bound for  $\mathbf{1}_Q(x)$  is the constant function with value 0. Therefore  $J^+(f) = 1 \times \lambda([0,T]) = T$  and  $J^-(f) = 0 \times \lambda([0,T]) = 0$ , which shows that  $\mathbf{1}_Q(x)$  is not Riemann integrable.

(d) For the Lebesgue integral we have

$$I(f) = \int_0^T f(x) dx = \int_0^T \mathbf{1}_Q(x) dx = \sum_{q \in [0,T] \cap \mathbb{Q}} \lambda(\{q\}) = 0$$

since the Lebesgue measure assigns zero mass  $\lambda(\{q\}) = 0$  to the singletons and  $\mathbb{Q}$  is countable.

5. (a) Prove Chebychev inequality: for a random variable X with  $X(\omega) \ge 0$  P-almost surely,

$$\mathbb{P}(X > t) \le \frac{E_{\mathbb{P}}(X)}{t} \quad \forall t > 0$$

Hint Note that

$$0 \le t \ \mathbf{1}(X(\omega) > t) \le X(\omega) \ .$$

(b) Prove Chentsov inequality

$$\mathbb{P}(X > t) \le \inf_{\theta > 0} \left\{ \exp(-\theta t) E_{\mathbb{P}}(\exp(\theta X)) \right\}$$

Hint: for any  $\theta > 0$ ,  $X > t \iff \exp(\theta X) > \exp(\theta t)$ .

**Solution** Since the expectation is a positive operator, the inequality is preserved after taking expectation:

$$0 \le tP(X > t) \le E_P(X)$$

Let  $Y = \exp(\theta X)$ . Since the map  $x \mapsto y = e^{\theta x}$  is strictly increasing when  $\theta > 0$ , seuraa  $\forall \theta > 0$ ,

$$P(X > t) = P(Y > \exp(\theta X)) \le E_P(\exp(\theta X))e^{-\theta t}$$
.

(c) Consider a random variable  $N(\omega)$  with Poisson( $\lambda$ ) distribution, where  $\lambda > 0$  is the parameter and

$$\mathbb{P}_{\lambda}(N=k) = \exp(-\lambda)\frac{\lambda^{k}}{k!} \quad k \in \mathbb{N} = \{0, 1, 2, \dots\}$$

(d) Knowing that  $E(\exp(\theta N)) = \exp(\lambda(e^{\theta} - 1))$ , (computed in the exercise sheet n.5) use Chentsov inequality to bound from above the probability  $\mathbb{P}_{\lambda}(N > t)$ , for t > 0.

$$P_{\lambda}(X > t) \le \inf_{\theta > 0} \left\{ \exp(-\theta t) E\left(\exp(\theta X)\right) \right\} = \inf_{\theta > 0} \left\{ \exp\left(\lambda(e^{\theta} - 1) - \theta t\right) \right\}$$

since at the minimum point  $t^\ast$ 

$$\frac{\partial}{\partial \theta} \left( \lambda(e^{\theta} - 1) - \theta t \right) = \lambda e^{\theta} - t = 0$$

with

$$\frac{\partial^2}{\partial \theta^2} \left( \lambda(e^{\theta} - 1) - \theta t \right) = \lambda e^{\theta} > 0,$$

it follows that the function  $\theta \mapsto (\lambda(e^{\theta} - 1) - \theta t)$  is convex and the minimum is achieved at  $\theta^* = \log(t) - \log(\lambda)$ , and we get the Chentsov upper bound as

$$P_{\theta}(X > t) \le \exp(t - \lambda) \left(\frac{\lambda}{t}\right)^{t}$$