

Random Walk, Martingales and Markov processes

1 Introduction

Evans discusses conditional expectations in §H of chapter 2 of his lecture notes [1]. The same topics can be found in § 2.1 of [4] where form the definitions of section 2 are taken. The definition of martingale follows instead § I of chapter 2 of [1]. A nice mathematical presentation of martingales in the case of countable state space is given in §11 of chapter 1 of [3]. The solution of the master equation for the Poisson process can be also found in §3.8.3 of [2].

2 Some definitions

Definition 2.1. Let $(\Omega, \mathcal{F}_n, \mathbb{P})$ be a probability space. A (discrete time) filtration is an increasing sequence $\mathcal{F} := \{\mathcal{F}_k\}_{k=0}^n$ of σ -algebras $\mathcal{F}_0 \subseteq \mathcal{F}_1 \subseteq \dots \mathcal{F}_n$. The quadruple $(\Omega, \mathcal{F}_n, \mathcal{F}, \mathbb{P})$ is called a filtered probability space

Definition 2.2. A stochastic process is just a collection of random variables $\{\xi_t, t \geq 0\}$, indexed by a time parameter t discrete or continuous.

Definition 2.3. Let $(\Omega, \mathcal{F}_n, \mathcal{F}, \mathbb{P})$ be a filtered probability space. A stochastic process $\{\xi_i\}_{i=1}^n$ is called \mathcal{F}_n -adapted ξ_n if is \mathcal{F}_n -measurable for every n , and is called \mathcal{F}_n -predictable if ξ_n is \mathcal{F}_{n-1} -measurable for every n .

Definition 2.4. Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and $\{\xi_n\}_n$ be a stochastic process. The filtration generated by $\{\xi_n\}_n$ is defined as $\mathcal{F}_n^\xi = \sigma(\xi_1, \dots, \xi_n)$ and the process is \mathcal{F}_n^ξ -adapted by construction.

3 Conditional Expectation, Heuristics

Let ξ an integrable random variable

$$\mathbb{E} \|\xi\| < \infty \quad (3.1)$$

on the probability space $(\Omega, \mathcal{F}, \mathbb{P})$. Let then \mathcal{G} a finite dimensional σ -algebra generated by the atomic decomposition (or partition) $\{A_i\}_{i=1}^n$ of Ω i.e.

$$\Omega = \bigcup_{i=1}^n A_i \quad (3.2)$$

In such a case we define the *conditional expectation* of ξ with respect to \mathcal{G} as the random variable

$$\mathbb{E}\{\xi|\mathcal{G}\} = \sum_{i=1}^n \frac{\mathbb{E}\{\xi\chi_{A_i}\}}{\mathbb{P}(A_i)} \chi_{A_i}(\omega) \quad (3.3)$$

where $\omega \in \Omega$ and

$$\chi_A(\omega) = \begin{cases} 1 & \text{if } \omega \in A \\ 0 & \text{if } \omega \notin A \end{cases} \quad (3.4)$$

Note that if the discrete random variable $\chi_A(\omega)$ is independent of ξ the definition implies immediately

$$E \{ \xi | \mathcal{G} \} = E \xi \sum_{i=1}^n \frac{E \{ \chi_{A_i} \}}{P(A_i)} \chi_{A_i}(\omega) \equiv E \xi \sum_{i=1}^n \chi_{A_i}(\omega) \quad (3.5)$$

whence

$$E \{ \xi | \mathcal{G} \} = E \xi \quad (3.6)$$

by virtue of

$$\sum_{i=1}^n \chi_{A_i}(\omega) = \chi_{\Omega}(\omega) = 1 \quad (3.7)$$

The latter equation just states that $\{A_i\}_{i=1}^n$ is a partition of Ω and that $\chi_{\Omega}(\omega)$ reduces to the trivial random variable equal to the unity whenever it is sampled i.e. $\forall \omega \in \Omega$.

Definition 3.1. Let (Ω, \mathcal{F}, P) be a probability space and suppose \mathcal{F}' is a σ -algebra, $\mathcal{F}' \subseteq \mathcal{F}$. If

$$\xi: \omega \mapsto \mathbb{R}^d \quad (3.8)$$

is an integrable random variable, we define

$$\xi' := E \{ \xi | \mathcal{F}' \} \quad (3.9)$$

to be any random variable on Ω such that

i ξ' is \mathcal{F}' -measurable;

ii For all $F' \in \mathcal{F}'$ the identity

$$E_P \{ \chi_{F'} \xi \} = E_P \{ \chi_{F'} \xi' \} \quad (3.10)$$

holds true.

If $\mathcal{F}' = \sigma(\eta)$ i.e. is the σ -algebra generated by η we will write

$$\xi' := E \{ \xi | \mathcal{F}' \} \equiv E \{ \xi | \eta \} \quad (3.11)$$

For more details please read § H of chapter 2 and/or § 2.1 of [4].

4 Martingales

Definition 4.1. Let $(\Omega, \mathcal{F}_n, \mathcal{F}, P)$ be a filtered probability space and $\{\xi_i\}_{i=1}^n$ a \mathcal{F}_n -adapted stochastic process such that

$$E |\xi_i| < \infty \quad \forall i \quad (4.1)$$

If

$$\xi_k = E \{ \xi_j | \mathcal{F}_k \} \quad a.s \quad \forall j \geq k$$

holds true we say that $\{\xi_i\}_{i=1}^n$ is a (discrete) martingale.

More generally we will consider stochastic processes $\{\xi_t, t \in T\}$ with T being a subset or coinciding with either \mathbb{R}_+ or \mathbb{N} . Furthermore we can posit that for any fixed $t \in T$ the random variable ξ_t takes values in a *state space* \mathbb{S} which may be finite (as for the random walk), countable or (a subset of) \mathbb{R}^d . In such a case the general definition is

Definition 4.2. Let \mathcal{F}_t be a filtration of the probability space $(\Omega, \mathcal{F}_t, P)$ and let $\{\xi_t, t \in T\}$ an \mathbb{S} -valued stochastic process adapted to \mathcal{F}_t satisfying

$$E \|\xi_t\| < \infty \quad (4.2)$$

for all $t \in T$. We say that $\{\xi_t, t \in T\}$ is a \mathcal{F}_t -martingale if

$$E \{\xi_t | \mathcal{F}_s\} = \xi_s \quad \forall t \geq s \in T \quad (4.3)$$

If instead

$$E \{\xi_t | \mathcal{F}_s\} \leq \xi_s \quad \forall t \geq s \in T \quad (4.4)$$

we say that $\{\xi_t, t \in T\}$ is a \mathcal{F}_t -super-martingale. Finally, if

$$E \{\xi_t | \mathcal{F}_s\} \geq \xi_s \quad \forall t \geq s \in T \quad (4.5)$$

we say that $\{\xi_t, t \in T\}$ is a \mathcal{F}_t -sub-martingale.

An important consequence of the martingale property is the conservation of the expectation value. Namely we must have

$$EE \{\xi_t | \mathcal{F}_s\} = E\xi_s \quad (4.6)$$

but also from the definition of conditional expectation

$$EE \{\xi_t | \mathcal{F}_s\} = E\xi_t \quad (4.7)$$

Hence for any t, s

$$E\xi_t = E\xi_s \quad (4.8)$$

5 Random Walk as Martingale

We defined the random walk as

$$\Xi_n = \sum_{i=1}^n \xi_i \quad (5.1)$$

with $\{\xi_i\}_{i=1}^N$ i.i.d. random variables with $\xi_i \stackrel{d}{=} \xi$ for all i . Furthermore

$$\xi: \Omega \mapsto \{-x, x\} \quad (5.2)$$

then

$$|\Xi_n| \leq n \quad \forall n = 1, \dots, N \quad (5.3)$$

We have

$$E \{\Xi_n | \Xi_n\} = \Xi_n \quad (5.4)$$

by definition of conditional expectation. We have

$$\mathbb{E} \{ \Xi_n | \Xi_{n-1} \} = \mathbb{E} \{ \Xi_{n-1} + \xi_n | \Xi_{n-1} \} = \mathbb{E} \{ \Xi_{n-1} | \Xi_{n-1} \} + \mathbb{E} \{ \xi_n | \Xi_{n-1} \} = \Xi_{n-1} + \mathbb{E} \{ \xi_n | \Xi_{n-1} \} \quad (5.5)$$

By definition of conditional expectation

$$\mathbb{E} \{ \xi_n | \Xi_{n-1} \} = \mathbb{E} \xi_n = (2p - 1) x \quad (5.6)$$

if $\mathbb{P}(\xi = x) = p$ as ξ_n is independent of Ξ_{n-1} . Repeating for arbitrary $k \leq n$

$$\mathbb{E} \{ \Xi_n | \Xi_k \} = \Xi_k + \sum_{i=k+1}^n \mathbb{E} \{ \xi_i | \Xi_k \} = \Xi_k + (n - k) (2p - 1) x \quad (5.7)$$

We verified that $\{ \Xi_n, 1 \leq n \leq N \}$ is

- a sub-martingale if $p > 1/2$;
- a martingale if $p = 1/2$;
- a super-martingale if $p < 1/2$.

From Ξ_n it is always possible to construct a martingale by subtracting its *compensator*:

$$\tilde{\Xi}_n = \Xi_n - A_n \quad (5.8)$$

In the case of the random walk

$$A_n = \sum_{i=1}^n \mathbb{E} \xi_i = n \mathbb{E} \xi = n (2p - 1) x \quad (5.9)$$

It is straightforward to verify that

$$\tilde{\Xi}_n = \sum_{i=1}^n \tilde{\xi}_i \quad (5.10)$$

is specified by the sum of i.i.d. random variables with zero average. Hence

$$\mathbb{E} \{ \tilde{\Xi}_n | \tilde{\Xi}_k \} = \tilde{\Xi}_k + \sum_{i=k+1}^n \mathbb{E} \{ \tilde{\xi}_i | \tilde{\Xi}_k \} = \tilde{\Xi}_k \quad (5.11)$$

which proves that $\{ \Xi_n, 1 \leq n \leq N \}$ is a martingale.

6 Markov process

Let us consider a stochastic process $\{ \xi_n, n \in \mathbb{N} \}$ valued to a countable state space $\mathbb{S} \subseteq \mathbb{Z}$:

$$\xi_n: \Omega \times \mathbb{N} \mapsto \mathbb{S} \quad (6.1)$$

We suppose that the evolution law for its probability distribution generalizes the form we found for the random walk

$$P_{n+1}(i) = \sum_{k \in \mathbb{S}} P(i, n+1 | k, n) P_n(k) \quad (6.2)$$

or equivalently

$$P_{n+1}(i) - P_n(i) = \sum_{k \in \mathbb{S}} [P(i, n+1|k, n) - \delta_{ik}] P_n(k) \quad (6.3)$$

Using the normalization condition

$$\sum_{k \in \mathbb{S}} P(k, n+1|i, n) = 1 \quad (6.4)$$

we can couch (6.3) into the form

$$P_{n+1}(i) - P_n(i) = \sum_{k \in \mathbb{S}} [P(i, n+1|k, n) P_n(k) - P(k, n+1|i, n) P_n(i)] \quad (6.5)$$

The master equation (6.2) states that at any time step we can reconstruct the probability of the stochastic process at the ensuing step if we know its “present” distribution. A more pictorial description is that the “future” depends only upon the “present” but not upon the “past”. Such a property is the distinguishing feature of Markov processes.

7 Continuous limit

Up to now we considered a unit time step. We may instead introduce a time unit τ and rescale probabilities

$$P_n(m) = \tilde{P}_{n\tau}(m) \quad (7.1)$$

The aim is to study the limit

$$\tau \downarrow 0 \quad \& \quad t = n\tau \in \mathbb{R}_+ \quad (7.2)$$

After rescaling we couch (6.5) into the form

$$\tilde{P}_{t+\tau}(i) - \tilde{P}_t(i) = \sum_{k \in \mathbb{S}} [\tilde{P}(i, t+\tau|k, t) \tilde{P}_t(k) - \tilde{P}(k, t+\tau|i, t) \tilde{P}_t(i)] \quad (7.3)$$

The expansion in Taylor series

$$\tilde{P}_{t+\tau}(i) = \tilde{P}_t(i) + \tau \partial_t \tilde{P}_t(i) + O(\tau^2) \quad (7.4)$$

and

$$\tilde{P}(i, t+\tau|k, t) = \delta_{ik} + \tau K_t(i|k) + O(\tau^2) \quad (7.5)$$

yields

$$\partial_t \tilde{P}_t(i) = \sum_{k \in \mathbb{S}} [K_t(i|k) \tilde{P}_t(k) - K_t(k|i) \tilde{P}_t(i)] + O(\tau) \quad (7.6)$$

Thus in the limit $\tau \downarrow 0$ we are left with

$$\partial_t \tilde{P}_t(i) = \sum_{k \in \mathbb{S}} [K_t(i|k) \tilde{P}_t(k) - K_t(k|i) \tilde{P}_t(i)] \quad (7.7)$$

Probability conservation now requires

$$\sum_{i \in \mathbb{S}} \partial_t \tilde{P}_t(i) = \partial_t \sum_{i \in \mathbb{S}} \tilde{P}_t(i) = 0 \quad (7.8)$$

entailing

$$\sum_{i,k \in \mathbb{S}} [\mathbb{K}_t(i|k) \tilde{\mathbb{P}}_t(k) - \mathbb{K}_t(k|i) \tilde{\mathbb{P}}_t(i)] = 0 \quad (7.9)$$

which is satisfied identically. Two observations are in order

- the diagonal component $\mathbb{K}_t(i|i)$ of the transition rate $\mathbb{K}_t(\cdot)$ does not contribute to (7.7).
- The condition

$$\sum_{k \in \mathbb{S}} \mathbb{K}_t(k|i) = 0 \quad (7.10)$$

is a sufficient condition for (7.9) to hold true. It also guarantees to leading order in $O(\tau)$ that

$$1 = \sum_{i \in \mathbb{S}} \tilde{\mathbb{P}}(i, t + \tau | k, t) = \sum_{i \in \mathbb{S}} [\delta_{i,k} + \tau \mathbb{K}_t(i|k) + O(\tau^2)] = 1 + O(\tau^2) \quad (7.11)$$

By virtue of the first observation, it is not restrictive to assume that (7.10) always holds true. In such a case we can write

$$\partial_t P_t(i) = \sum_{k \in \mathbb{S}} [\mathbb{K}_t(i|k) P_t(k) \quad (7.12a)$$

$$\sum_{k \in \mathbb{S}} \mathbb{K}_t(k|i) = 0 \quad (7.12b)$$

8 Poisson process

We now make a special choice for the transition rates in (7.12) and set for some $\gamma \in \mathbb{R}_+$

$$\mathbb{K}_t(i|k) = \gamma \delta_{i,k+1} - \gamma \delta_{k,i} \quad (8.1)$$

The resulting equation is

$$\partial_t P(i, t) = \gamma P(i-1, t) - \gamma P(i, t)$$

This is the evolution for a process that can make (or not make) jumps only towards the right of its current position. If we assume that the initial distribution

$$P(i, 0) = P_o(i)$$

has support on \mathbb{N} then the process will stay there for any further time. The equation can be solved exactly by computing the characteristic function

$$\tilde{\mathbb{P}}(u, t) := \sum_{k=0}^{\infty} e^{\lambda k u} P(k, t)$$

Namely, it is straightforward to see that $\tilde{\mathbb{P}}(u, t)$ satisfies:

$$\partial_t \tilde{\mathbb{P}}(u, t) = \gamma (e^{\lambda u} - 1) \tilde{\mathbb{P}}(u, t)$$

The solution for the initial condition $\check{P}(u, 0) = \check{P}_o(u)$

$$\check{P}(u, t) = e^{\gamma t (e^u - 1)} \check{P}_o(u)$$

If we specialize for an initial condition

$$P_o(i) = \delta_{i0}$$

(i.e. we assume that the process starts from the origin) we obtain

$$\check{P}(u, t) = e^{\gamma t (e^u - 1)}$$

In order to infer the probability distribution associated to the characteristic function we can write

$$\check{P}(u, t) = e^{\gamma t e^u} e^{-\gamma t} = e^{-\gamma t} \sum_{j=0}^{\infty} \frac{(\gamma t)^j}{\Gamma(j+1)} e^{u j}$$

which implies that $\check{P}(u, t)$ is the characteristic function of the **Poisson process**, with probability distribution:

$$P(j, t) = \frac{(\gamma t)^j}{\Gamma(j+1)} e^{-\gamma t}$$

References

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