

# 1 Introduction

These notes shortly recall some basic concepts in classical probability. The material presented here is more leisurely expounded in a physics style language in chapter two of [2]. For a more mathematics style presentation see chapter one of [3] or chapter two of [1].

## 2 Measure theoretic definitions

Let  $\Omega$  a non-empty set.

**Definition 2.1** ( *$\sigma$ -algebra*). A  $\sigma$ -algebra is a collection  $\mathcal{F}$  of subsets of  $\Omega$  with these properties

1.  $\emptyset, \Omega \in \mathcal{F}$ .
2. if  $F \in \mathcal{F}$  then  $F^c \in \mathcal{F}$  for  $F^c := \Omega \setminus F$  the complement of  $F$ .
3. if  $\{F_k\}_{k=1}^{\infty} \in \mathcal{F}$  then

$$\bigcap_{k=1}^{\infty} F_k, \bigcup_{k=1}^{\infty} F_k \in \mathcal{F}$$

**Definition 2.2** (*Probability measure*). Let  $\mathcal{F}$  be a  $\sigma$ -algebra of subsets of  $\Omega$ . We call

$$P : \mathcal{F} \rightarrow [0, 1]$$

a probability measure provided:

1.  $P(\emptyset) = 0, P(\Omega) = 1$
2. if  $\{F_k\}_{k=1}^{\infty}$  then

$$P(\bigcup_{k=1}^{\infty} F_k) \leq \sum_{k=1}^{\infty} P(F_k)$$

3. if  $\{F_k\}_{k=1}^{\infty}$  are **disjoint sets**

$$P(\bigcup_{k=1}^{\infty} F_k) = \sum_{k=1}^{\infty} P(F_k) \tag{2.1}$$

It follows that if  $F_1, F_2 \in \mathcal{F}$

$$F_1 \subset F_2 \quad \Rightarrow \quad P(F_1) \leq P(F_2)$$

**Definition 2.3** (*Borel  $\sigma$ -algebra*). The smallest  $\sigma$ -algebra containing all the **open** subsets of  $\mathbb{R}^d$  is called the Borel  $\sigma$ -algebra, denoted by  $\mathcal{B}$

The **Borel subsets** of  $\mathbb{R}^d$  i.e. the content of  $\mathcal{B}$  may be thought as the collection of all the well-behaved subsets of  $\mathbb{R}^d$  for which Lebesgue measure theory applies.

### 3 Probability Space

**Definition 3.1** (*Probability space*). A triple

$$(\Omega, \mathcal{F}, P)$$

is called a probability space provided

1.  $\Omega$  is any set
  2.  $\mathcal{F}$  is a  $\sigma$ -algebra of subsets of  $\Omega$
  3.  $P$  is a probability measure on  $\mathcal{F}$
- Points  $\omega \in \Omega$  are sample (outcome) points.
  - A set  $F \in \mathcal{F}$  is called an event.
  - $P(F)$  is the probability of the event  $F$ .
  - A property which holds true but for events of probability zero is said to hold almost surely (usually abbreviated "a.s.").

**Example 3.1** (*Single unbiased coin tossing*). :

- outcomes: head, tail
- $\Omega = \{head, tail\}$ .
- $\sigma$ -algebra  $\mathcal{F}$ : it comprises  $|\mathcal{F}| = 2^{|\Omega|} = 4$  events
  - 1 T=tail
  - 2 H=head
  - 3  $\emptyset$ =neither head nor tail
  - 4  $T \vee H$ =head or tail
- Probability measure:

$$P(T) = P(H) = \frac{1}{2} \quad \& \quad P(\emptyset) = 0 \quad \& \quad P(T \vee H) = 1 \quad (3.1)$$

**Example 3.2** (*Uniform distribution*). :

- $\Omega = [0, 1]$ .
- $\mathcal{F}$ : the  $\sigma$ -algebra of all Borel subsets of  $[0, 1]$ .
- $P$ : the Lebesgue measure on  $[0, 1]$ . (Note: as  $0 \cup 1$  has zero measure  $[0, 1] \sim (0, 1)$ .)

**Definition 3.2** (*Probability density on  $\mathbb{R}^d$* ). Let  $p$  be a **non-negative**, integrable function, such that

$$\int_{\mathbb{R}^d} d^d x p(\mathbf{x}) = 1 \quad (3.2)$$

then to each  $B \in \mathcal{B}$  (Borel  $\sigma$ -algebra) is possible to associate a probability

$$P(B) = \int_B d^d x p(\mathbf{x}) \quad (3.3)$$

so that  $(\mathbb{R}^d, \mathcal{B}, P)$  is a probability space. The function  $p$  is called the density of the probability  $P$ .

**Example 3.3** (*Gaussian distribution*). The function

$$g_{\bar{x}\sigma} : \mathbb{R} \rightarrow \mathbb{R}_+$$

$$g_{\bar{x}\sigma}(x) = \frac{e^{-\frac{(x-\bar{x})^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}} \quad (3.4)$$

is a probability density on  $(\mathbb{R}^d, \mathcal{B}, P)$ .

## 4 Random variables

**Definition 4.1** ( *$\mathbb{R}^d$ -valued random variable*). Let  $(\Omega, \mathcal{F}, P)$  be a probability space. A mapping

$$\xi : \Omega \rightarrow \mathbb{R}^d$$

is called an  $d$ -dimensional **random variable** if for each  $B \in \mathcal{B}$  ( $\mathcal{B}$  is the Borel  $\sigma$ -algebra over  $\mathbb{R}^d$ ) one has

$$\xi^{-1}(B) \in \mathcal{F}$$

i.e. if  $\xi$  is  $\mathcal{F}$ -measurable.

The definition associates to each event a Borel subset. More generally if  $\mathbb{S}$  is a set and  $\mathcal{S}$  a  $\sigma$ -algebra on it such that the pair  $(\mathbb{S}, \mathcal{S})$  is a measurable space then we have

**Definition 4.2** ( *$\mathbb{S}$ -valued random variable*). Let  $(\Omega, \mathcal{F}, P)$  be a probability space. A mapping

$$\xi : \Omega \rightarrow \mathbb{S}$$

is an  $\mathbb{S}$ -valued random variable if for each  $S \in \mathcal{S}$  one has

$$\xi^{-1}(S) \in \mathcal{F}$$

See [3] 1.3 for further details.

**Example 4.1** (*Indicator function*). Let  $F \in \mathcal{F}$ . The indicator function of  $F$  is

$$\chi_F(\omega) = \begin{cases} 1 & \text{if } \omega \in F \\ 0 & \text{if } \omega \notin F \end{cases}$$

**Example 4.2** (*Simple function*). Let  $\{F_i\}_{i=1}^m \in \mathcal{F}$  are disjoint (i.e.  $F_i \cap F_j = \emptyset$ ) and form a partition of  $\Omega$  (i.e.  $\cup_{i=1}^m F_i = \Omega$ ) and  $\{x_i\}_{i=1}^m \in \mathbb{R}$  then

$$\xi = \sum_{i=1}^m x_i \chi_{F_i}(\omega)$$

is a random variable, called a **simple function**.

**Lemma 4.1.** *Let*

$$\xi(\omega) : \Omega \rightarrow \mathbb{R}^d$$

*be a random variable. Then*

$$\mathcal{F}(\xi) = \{\xi^{-1}(B) \mid B \in \mathcal{B}\}$$

*is a  $\sigma$ -algebra, called the  $\sigma$ -algebra generated by  $\xi$ . This is the smallest sub  $\sigma$ -algebra of  $\mathcal{F}$  with respect to which  $\xi$  is measurable.*

*Proof.* It suffices to verify that  $\mathcal{F}(\xi)$  is a  $\sigma$ -algebra. □

**Remark 4.1 (Meaning of measurability).** : The  $\sigma$ -algebra  $\mathcal{F}(\xi)$  encodes all the information described by the random variable  $\xi$ . This means that if  $\zeta$  is a second random variable, the statement

- $\zeta = f(\xi)$  for some mapping  $f$  implies that  $\zeta$  is  $\mathcal{F}(\xi)$ -measurable.
- $\zeta$  is  $\mathcal{F}(\xi)$ -measurable, implies that there exists a mapping  $f$  such that  $\zeta = f(\xi)$ .

## 5 Expectation values

Expectation values of generic random variables are defined following the same steps taken to define the Lebesgue integral of measurable functions. Let  $(\Omega, \mathcal{F}, P)$  a probability space and  $\xi$  a simple 1-dimensional random variable

$$\xi = \sum_{i=1}^n x_i \chi_{F_i}$$

**Definition 5.1 (Expectation value (integral) of a simple random variable).**

$$E\xi := \int_{\Omega} dP \xi = \sum_{i=1}^n x_i P(F_i)$$

**Definition 5.2 (Expectation value (integral) of a non-negative random variable  $\eta$ ).** For

$$\eta : \Omega \rightarrow \mathbb{R}_+$$

*we define*

$$E\eta \equiv \int_{\Omega} dP \eta := \sup_{\substack{\xi \leq \eta \\ \xi = \text{simple}}} \int_{\Omega} dP \xi$$

**Definition 5.3 (Expectation value a random variable  $\eta$ ).** For

$$\eta : \Omega \rightarrow \mathbb{R}$$

*we define*

$$\eta_+ := \max\{\eta, 0\} \quad \& \quad \eta_- := \max\{-\eta, 0\}$$

*If*

$$\min\{E\eta_+, E\eta_-\} < \infty$$

we define the expectation variable of

$$\eta \equiv \eta_+ - \eta_-$$

as

$$\int_{\Omega} dP \eta := \int_{\Omega} dP \eta_+ - \int_{\Omega} dP \eta_-$$

With these definitions all the standard rules of Lebesgue integrals apply to expectation values.

## 6 Moments of a random variable

**Definition 6.1 (Distribution function).** The distribution function of a random variable  $\xi : \Omega \rightarrow \mathbb{R}^d$  is the function

$$\tilde{P}_{\xi} : \mathbb{R}^d \rightarrow \mathbb{R}_+$$

such that

$$\tilde{P}_{\xi}(\mathbf{x}) = P_{\xi}(\xi_1 \leq x_1, \dots, \xi_d \leq x_d)$$

**Definition 6.2 (PDF of a random variable).** Let  $\xi : \Omega \rightarrow \mathbb{R}^d$  be a random variable and  $\tilde{P}_{\xi}$  its distribution function. If there exists a **non-negative, integrable function**

$$p : \mathbb{R}^d \rightarrow \mathbb{R}_+$$

such that

$$\tilde{P}_{\xi}(\mathbf{x}) = \prod_{i=1}^d \int_{-\infty}^{x_i} dy_i p_{\xi}(\mathbf{y})$$

then  $p_{\xi}$  specifies the probability density function of  $\xi$  (PDF).

**Lemma 6.1.** Let

$$\xi : \Omega \rightarrow \mathbb{R}^d$$

be a random variable, with statistics described by PDF  $p_{\xi}$ . Suppose

$$f : \mathbb{R}^d \rightarrow \mathbb{R}$$

and

$$y = f(\mathbf{x})$$

Then

$$\mathbb{E}f(\xi) = \int d^d x p_{\xi}(\mathbf{x}) f(\mathbf{x})$$

*Proof.* Suppose first  $f$  is a **simple** function on  $\mathbb{R}^d$ . Then

$$\mathbb{E}f(\xi) = \sum_{i=1}^n f_i \int \chi_{B_i}(\xi) dP = \sum_{i=1}^n f_i P(B_i) = \sum_{i=1}^n \int_{B_i} p_{\xi}(\mathbf{x}) f(\mathbf{x})$$

Consequently the formula holds for all simple functions  $g$  and, by approximation, it holds therefore for general functions  $g$ .  $\square$

In particular for a  $\mathbb{R}^d$ -valued random variable we have

$$\mathbb{E}\boldsymbol{\xi} = \int d^d x p_{\boldsymbol{\xi}}(\mathbf{x}) \mathbf{x} \quad \text{average or mean value}$$

meaning that if  $\{\mathbf{e}_i\}_{i=1}^d$  is the collection of the unit vectors spanning the canonical basis of  $\mathbb{R}^d$  i.e.

$$\mathbf{e}_1 = \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix} \quad \mathbf{e}_2 = \begin{bmatrix} 0 \\ 1 \\ \vdots \\ 0 \end{bmatrix} \quad \text{etc.} \quad (6.1)$$

then

$$\mathbf{e}_i \cdot \mathbb{E}\boldsymbol{\xi} = \mathbb{E}\mathbf{e}_i \cdot \boldsymbol{\xi} \equiv \mathbb{E}\xi^i = \int d^d x p_{\boldsymbol{\xi}}(\mathbf{x}) x^i \quad (6.2)$$

Similarly

$$\mathbb{E}(\boldsymbol{\xi} - \mathbb{E}\boldsymbol{\xi}) \otimes (\boldsymbol{\xi} - \mathbb{E}\boldsymbol{\xi}) = \int d^d x p_{\boldsymbol{\xi}}(\mathbf{x}) (\mathbf{x} - \mathbb{E}\boldsymbol{\xi}) \otimes (\mathbf{x} - \mathbb{E}\boldsymbol{\xi}) \quad \text{(co-)variance tensor}$$

stands for the rank two tensor with components

$$\begin{aligned} \mathbb{E}(\xi^i - \mathbb{E}\xi^i)(\xi^j - \mathbb{E}\xi^j) &= \int d^d x p_{\boldsymbol{\xi}}(\mathbf{x}) (x^i - \mathbb{E}x^i)(x^j - \mathbb{E}x^j) \\ &= \int d^d x p_{\boldsymbol{\xi}}(\mathbf{x}) (x^i x^j - x^j \mathbb{E}x^i - x^i \mathbb{E}x^j) + \mathbb{E}x^i \mathbb{E}x^j = \int d^d x p_{\boldsymbol{\xi}}(\mathbf{x}) x^i x^j - \mathbb{E}x^i \mathbb{E}x^j \end{aligned} \quad (6.3)$$

Note that for any (non-random) vector  $\mathbf{v}$  the inequality

$$\mathbf{v} \cdot \mathbb{E}(\boldsymbol{\xi} - \mathbb{E}\boldsymbol{\xi}) \otimes (\boldsymbol{\xi} - \mathbb{E}\boldsymbol{\xi}) \cdot \mathbf{v} \equiv \mathbb{E}[(\boldsymbol{\xi} - \mathbb{E}\boldsymbol{\xi}) \cdot \mathbf{v}]^2 \geq 0 \quad (6.4)$$

holds true. The covariance tensor also referred to as correlation tensor is therefore positive definite and in particular strictly positive definite for any random variable  $\boldsymbol{\xi}$  whose probability distribution does not degenerate on a single deterministic value.

**Definition 6.3** (*Moments of a random variable*). Let

$$\xi : \Omega \rightarrow \mathbb{R}$$

we call the expectation value of the  $n$ -th power of  $\xi$

$$\mathbb{E}\xi^n = \int_{\Omega} dP \xi^n$$

the moment of order  $n$  of  $\xi$ .

The lower order moments are those most recurrent in applications and as such are given special names such as the average and the variance.

**Example 6.1.** (*Average and variance of a Gaussian variable*)

- Average:

$$E \xi = \int_{\mathbb{R}} dx x g_{\bar{x}\sigma}(x) = \int_{\mathbb{R}} dx (\bar{x} + \sigma x) g_{01}(x)$$

As

$$g_{01}(x) = g_{01}(-x)$$

we find

$$E \xi = \bar{x}$$

- Variance

$$E (\xi - E \xi)^2 = \sigma^2 \int_{\mathbb{R}} dx x^2 g_{01}(x)$$

The remaining integral  $I$  can be evaluated for example using the identity

$$\int_{\mathbb{R}} dx x^2 g_{01}(x) = \left. \frac{d^2}{dj^2} Z(j) \right|_{j=0}$$

$$Z(j) := \int_{\mathbb{R}} dx g_{01}(x) e^{jx} = \int_{\mathbb{R}} dx \frac{e^{-\frac{(x-j)^2}{2}}}{\sqrt{2\pi}} e^{\frac{j^2}{2}}$$

Performing the change of variable  $x \mapsto x + j$  we can therefore write

$$Z(j) = e^{\frac{j^2}{2}} \int_{\mathbb{R}} dx \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} = e^{\frac{j^2}{2}}$$

Finally in order to prove that  $g_{01}(x)$  is indeed normalized to the unity we observe that

$$\left[ \int_{\mathbb{R}} dx \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} \right]^2 = \int_{\mathbb{R}^2} \prod_{i=1}^2 dx_i \frac{e^{-\frac{x_1^2+x_2^2}{2}}}{2\pi} = \int_0^\infty dr r e^{-\frac{r^2}{2}} = 1$$

The statistical properties of a Gaussian variable are therefore fully specified by its first two moments.

## 7 Characteristic function of a random variable

**Definition 7.1** (*Characteristic function*). Let

$$\xi : \Omega \rightarrow \mathbb{R}^d$$

the expectation value

$$\check{p}_\xi(\mathbf{q}) := E e^{i\mathbf{q}\cdot\xi}$$

is referred to as the characteristic function of the random variable

**Example 7.1** (*Characteristic function of a Gaussian random*). Let

$$\xi : \Omega \rightarrow \mathbb{R}$$

distributed with Gaussian PDF. The characteristic function is

$$\check{g}_{\bar{x},\sigma}(q) = \int dx e^{iqx} g_{\bar{x},\sigma}(x) = e^{iqx - \frac{\sigma^2 q^2}{2}}$$

For a Gaussian variable it is also true

$$\check{g}_{\bar{x},\sigma}(q) = \sum_{n=0}^{\infty} \frac{i^n q^n}{\Gamma(n+1)} \prec \xi^n \succ$$

having used the  $\Gamma$ -function representation of the factorial (see appendix A). If  $\bar{x} = 0$  odd moments vanish whilst for even moments we get into

$$\mathbb{E} \xi^{2n} = \sigma^{2n} \frac{\Gamma(2n+1)}{2^n \Gamma(n+1)} = (2n-1)!! \sigma^{2n} \quad (7.1)$$

**Formally** for a random variable  $\xi$  one can write

$$\mathbb{E} \xi^n = \frac{1}{i^n} \left. \frac{d^n}{dq^n} \check{p}_{\xi}(q) \right|_{q=0} \quad (7.2)$$

The relation is formal because it may be a relation between infinities.

**Example 7.2** (*Lorentz distribution*). Let

$$p : \mathbb{R} \rightarrow \mathbb{R}_+$$

be

$$p_{y\sigma}(x) = \frac{\sigma}{\pi \{(x-y)^2 + \sigma^2\}}$$

the Lorentz probability density so that  $(\mathbb{R}, \mathcal{B}, P_{y\sigma}(\mathcal{B}))$  a probability space. Note that

$$p_{0\sigma}(x) = p_{0\sigma}(-x)$$

Using a change of variable and it is straightforward to verify that

$$\int_{\mathbb{R}} dx x p_{y\sigma}(x) = y$$

however

$$\int_{\mathbb{R}} dx x^2 p_{y\sigma}(x) = \infty$$

The characteristic function can be computed using Cauchy theorem

$$\begin{aligned} \check{p}_{y\sigma}(q) &= e^{iqy} \int dx e^{iqx} p_{0\sigma}(x) \\ &= \frac{e^{iqy}}{2i\pi} \int_{\mathbb{R}} dx e^{iqx} \left\{ \frac{1}{x-i\sigma} - \frac{1}{x+i\sigma} \right\} = e^{iqy} \begin{cases} e^{-q\sigma} & \text{if } q > 0 \\ e^{q\sigma} & \text{if } q < 0 \end{cases} \end{aligned}$$

The characteristic function develops a cusp for  $q = 0$

$$\check{p}_{y\sigma}(q) = e^{iqy - \sigma|q|}$$



## Appendices

### A Gamma function

The  $\Gamma$  function for any  $x \in \mathbb{R}_+$  is specified by the integral

$$\Gamma(x) = \int_0^{\infty} \frac{dy}{y} y^x e^{-y}$$

For  $x \in \mathbb{N}$  the integral can be performed explicitly and it is equal to the factorial:

$$\Gamma(x) = (x - 1)! \quad x \in \mathbb{N}$$

For  $x \in \mathbb{R}_+$ , integration by parts yields the identity

$$\Gamma(x + 1) = \int_0^{\infty} \frac{dy}{y} y^{x+1} e^{-y} = x \Gamma(x)$$

which is trivially satisfied by factorials. For  $x \gg 1$  the value of the integral is approximated by *Laplace's stationary point* method

$$\Gamma(x + 1) \simeq e^{x(\ln x - 1)} \int_{\mathbb{R}} dy e^{-\frac{y^2}{2x}} = \sqrt{2\pi x} e^{x(\ln x - 1)} \quad x \gg 1 \quad (\text{A.1})$$

Such asymptotic estimation is usually referred to as *Stirling formula*.

### References

- [1] L. C. Evans. An introduction to stochastic differential equations. UC Berkeley, Lecture Notes, 2006.
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- [3] R. van Handel. Stochastic calculus and stochastic control. Lecture Notes, Caltech, 2007.