- Collect several definitions by using indexing and looping:
  - for(i in 1:K){ X[i] ~ dbin(theta[i],N[i]) }
  - Give K as fixed value in the data listing.
  - Every definition within "for(i in 1:K){ }" should have index i.
  - Can make several nested loops, for "x[i,j]" etc.
  - Can use nested indexing, for "x[y[i]]".
  - Can use arithmetics in indexing, for "x[i+10]"
  - Be careful to index correctly!

- What distributions and logical functions are available?
  - Check the list from manual/menu.
  - Pay attention to parameterization!
  - A very useful deterministic function: step()
  - What you define should be logically correct and computable in all situations.
     1/X should not become 1/0 with stochastic node X.

• Data formatting (every data variable should appear in the model)

```
list(x=4,
    y=c(3.5,7.2,9.1),
    z=structure(
    .Data=c(7,3,5,1,8,2),
    .Dim=c(2,3)) )
```

**Alternative format** 

```
z[,1] z[,2] z[,3]
7 3 5
1 8 2
END (Note: empty line after END)
```

 Irregular data can be coded using NA for "missing"

list( z=structure(

```
.Data=c(7,NA,NA,
```

```
9,6,3,
2,NA,5),
```

.Dim=c(3,3)))

BUGS would generate predictions for NAs. Alternatively, use auxiliary indexing: list(z=c(7,9,6,3,2,5),person=c(1,2,2,2,3,3))

 Transformations of data can be declared within model code
 yy <- log(y)</li>
 yy ~ dnorm(mu,tau)

Here y would be given in data.

• Can check your data values from 'info'  $\rightarrow$  'node info'  $\rightarrow$  'values'

- Diagnostic testing.
- Estimate a proportion.
- Compare proportions in two populations.
- Estimate a mean.
- Compare means in two populations.
- Linear & other regression.

- Diagnostic testing with additional data on sensitivity and specificity model{
- x[1] ~ dbin(q[1],N[1]) x[2] ~ dbin(q[2],N[2]) y ~ dbin(pr,M); pr <- p\*q[1]+(1-p)\*(1-q[2]) q[1] ~ dunif(0,1); q[2] ~ dunif(0,1); p ~ dunif(0,1)
- }
  list(x=c(45,28),N=c(50,30),M=100,y=10)

- Estimate and compare proportions model{
  - x ~ dbin(p,Nx); p ~ dunif(0,1)
  - y ~ dbin(q,Ny); q ~ dunif(0,1)
  - diff <- p q # could assess the difference
  - r <- p/q # could assess the ratio
  - pr <- step(diff) # an indicator variable</pre>

list(x=3,y=7,Nx=20,Ny=45)

}

Estimate and compare means

```
model{
for(i in 1:N){
 ahonen[i] ~ dnorm(m[1],tau[1])
janda[i] ~ dnorm(m[2],tau[2])
for(i in 1:2){
 m[i] ~ dnorm(0,0.0001)
tau[i] ~ dgamma(0.01,0.01)
 diff <- m[1]-m[2]
 pr <- step(diff)</pre>
```

 Data: scores of J Ahonen and J Janda from 8 ski jumping competitions (Four Hills tournament, 2006)

list( ahonen = c(299.7, 255.2, 281.7, 238.0, 270.9, 262.2, 255.4, 293.0), janda = c(238.7, 285.6, 287.1, 252.2, 262.6, 264.7, 263.2, 291.0), N=8)

- Tips: with very flat, vague priors, BUGS may generate bad starting values (better to assign inits yourself).
- Generate predictions by adding this line: pred.ahonen ~ dnorm(m[1],tau[1])

- In 2006, they got exactly the same total score!
- Assume you have the results after 7 competitions. Make a prediction for the total score. What's the probability that the difference is < 1 point?
- Set the 8th result as 'NA' in data, then run the following: model{

```
for(i in 1:N){
  ahonen[i] ~ dnorm(m[1],tau[1])
  janda[i] ~ dnorm(m[2],tau[2])
  }
  for(i in 1:2){
  m[i] ~ dnorm(0,0.0001)
  tau[i] ~ dgamma(0.01,0.01)
  }
  ahonen.total <- sum(ahonen[1:N])
  janda.total <- sum(janda[1:N])
   pr <- 1- step(abs(ahonen.total-janda.total)-1)
  }
</pre>
```

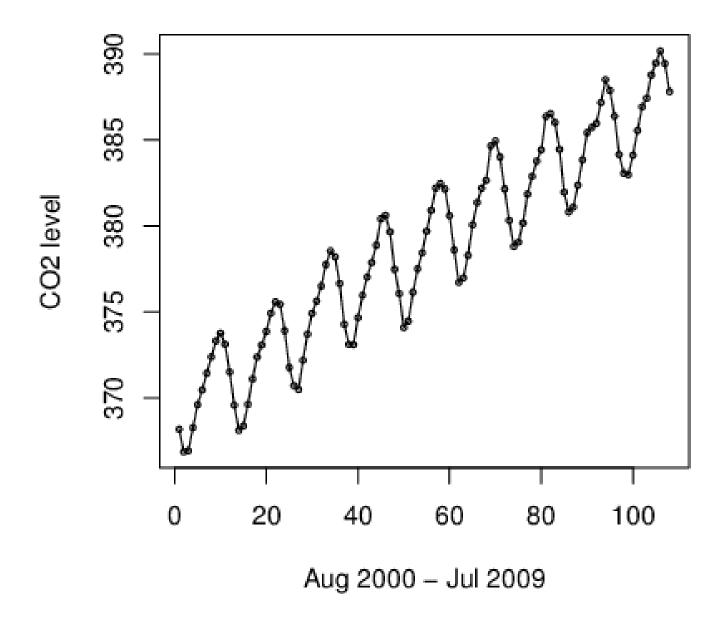
- Linear regression
  - York rainfall data: x= in November, y= in December

```
model{
                           List(y = c(41,52,18.7,55,40,29.2,51,17.6,46.6,57),
                          x = c(23.9,43.3,36.3,40.6,57,52.5,46.1,142,112.6,23.7))
for(i in 1:10){
y[i] ~ dnorm(mu[i],tau)
mu[i] <- beta[1] + beta[2]*x[i]
# mu[i] <- beta[1]+ beta[2]*(x[i]-mean(x[])) # standardized covariates</pre>
for(i in 1:2){
beta[i] \sim dnorm(0,0.001)
tau \sim dgamma(0.01, 0.01)
# prediction with given value xnew:
ynew ~ dnorm(munew,tau); munew <- beta[1] + beta[2]*xnew</pre>
# munew <-beta[1] + beta[2]*(xnew-mean(x[])) # standardized covariates</pre>
Interpretation of beta[1] in both cases? E(y | x=0) vs E(y | x=mean(x))
```

- Linear & nonlinear regression
  - Atmospheric CO<sub>2</sub>, monthly, Mauna Loa, Hawaii

list(N=120,x=c(368.18,366.87,366.94,368.27,369.62,370.47, 371.44,372.39,373.32,373.77,373.13,371.51,369.59,368.12, 368.38,369.64,371.11,372.38,373.08,373.87,374.93,375.58, 375.44,373.91,371.77,370.72,370.5,372.19,373.71,374.92, 375.63,376.51,377.75,378.54,378.21,376.65,374.28,373.12, 373.1,374.67,375.97,377.03,377.87,378.88,380.42,380.62, 379.66,377.48,376.07,374.1,374.47,376.15,377.51,378.43, 379.7,380.91,382.2,382.45,382.14,380.6,378.6,376.72, 376.98,378.29,380.07,381.36,382.19,382.65,384.65, 384.94,384.01,382.15,380.33,378.81,379.06,380.17, 381.85,382.88,383.77,384.42,386.36,386.53,386.01, 384.45,381.96,342, 385.72,385.96,387.18,388.5,387.88,386.38,384.15,

383.07,382.98,384.11,385.54,386.93,387.42,388.77, 389.46,390.18,389.43,387.81)



Linear and nonlinear terms

```
model{
tau \sim dgamma(0.01, 0.01);
for(i in 1:5)\{a[i] \sim dnorm(0,0.001)\}
for(i in 1:N){
month[i] <- i
x[i] ~ dnorm(mu[i],tau)
mu[i] <- a[1] + a[2]^{i} + a[3]^{sin}(2^{pi^{i}}/12) + a[4]^{cos}(2^{pi^{i}}/12)
pi <- 3.1415926
```

- Generalized linear model: Poisson
  - Number of lung cancer cases
  - Population counts
  - In age groups, in different cities, in 1968-1971.
  - Use the first age group in the first city as a reference, to compute age effects and city effects
  - $\log(\lambda_{\text{age,city}}) = \mu_0 + \alpha_{\text{age}} + \beta_{\text{city}}$ , with  $\alpha_{\text{age=1}} = \beta_{\text{city=1}} = 0$
  - $X_{age,city} \sim Poisson(4\lambda_{age,city} pop_{age,city})$

```
model{ # design matrix X could also be written beforehand in data
         # but it is here constructed from 'age' and 'city'.
         # The linear predictor can then be computed using inprod.
for(i in 1:24){
cases[i] ~ dpois(mu[i]); group[i] <- i
mu[i] <- pop[i]*4*lambda[i] # lambda = incidence per year
LA[i] <- lambda[i]/100000 # LA = inc. per 10^5 per year
log(lambda[i]) <- inprod(alpha[],X[i,])
X[i,1] <- 1
for(k in 2:6){X[i,k] <- equals(age[i],k) }</pre>
for(k in 2:4){X[i,k+5] <- equals(city[i],k)}
for(k in 1:9){ alpha[k] ~ dnorm(0,0.001)
A[k] <- exp(alpha[k]) }
```

# Tips

- Always think it as a DAG.
- Data variable has to correspond to a (fixed) stochastic "~"node in the model code, not "<-". The latter would make 'multiple deterministic definitions' error.
- Ddistr(?,?) ← Parameters, not expressions. Check parameterization !
- Test first with a small number of iterations how slow or fast it is.
- Give constants in data, not within code.
- Separate clearly what's data, what's model.
- Use comments # there are never too many!
- Collect definitions logically into groups (priors, likelihoods, predictions), easier to read.
- Transformations of data can be defined within code.
- Use indexing, and nested indexing.
- Avoid multiple definitions (e.g. within for-loops!) they are syntax errors.
- Break long expressions into short ones (avoid 'logical expression too complex' error)
- Pay attention to naming of parameters, variables. They should be meaningful at first sight. (or write good explanations in comment lines)
- Constants cannot be monitored, but can check them from node-info menu button.
- Sooner or later, it will be more convenient to run BUGS from R, try it.
- For the inbuilt convergence diagnostics, you should pick overdispersed starting values for at least 3 chains.
- Think of identifiability: is there sufficient data? Is something hanging completely from prior? It is deceptively easy to build castles in the clouds....
- Make use of inprod to avoid writing long expressions a[1]\*X[1]+a[2]\*X[2]+......