Calibrating on Principal Components in the presence of Multiple Auxiliary Variables for Nonresponse Adjustment

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Introduction-Motivation

- Auxiliary information plays a prominent role in successful estimation.
 Rizzo, Kalton and Brick (1996) note that, providing it is carefully chosen, the particular adjustment scheme used at the estimation stage is not that important.
- Increase of storage capacities and automatic data collection process may lead to data analysis of large set of auxiliary variables.
- The use of large sets of auxiliary variables may incur in constructing estimators with worse performances (inefficient than HT, Cardot, Goga and Shehzad, 2014)
- The Computational effort increases with the number of variables
- Common practice of usage of large sets of auxiliary variables is based on selection criteria of subsets deemed important auxiliary variables
- In multipurpose estimations it is less likely that specific variable become important to all study variables or at least all important study variables

Principal Components as an alternative to selection criteria

 Gives an optimal linear combination of all candidate auxiliary variables

 Produce a joint effect of the candidate variables on the variables of interest.

 Reduces the computational effort due to large number of auxiliary variables

Summary on Principal Components

• Let $\mathbf{X}_{(N \times P)}$ be our auxiliary data matrix, and suppose that

$$\mathbf{X}_{j(N\! imes\!1)}, \, \left\{ j\!=\!1\!,...,\!P\!
ight\}$$
 is rescaled to zero mean and unit variance, then

$$\mathbf{X}^t\mathbf{X}$$
 is the covariance matrix of \mathbf{X} . If $\left(\lambda_j,\mathbf{b}_j;j=1,...,P\right)$ are pairs

eigenvalue-eigenvector of
$$\mathbf{X}^t\mathbf{X}$$
 , then, the \mathbf{J}^{th} principal components

is given by
$$Z_j = \mathbf{b}_j^t \mathbf{X}$$
 all Z_j 's are uncorrelated.

• When $\mathbf{X} = \mathbf{X}_{(n \times P)}$, then, $\operatorname{cov}(\mathbf{X}) = \mathbf{X}^I D \mathbf{X}$ where $D = \operatorname{diag}\left\{d_1, \ldots, d_n\right\}, d_k = 1 / \pi_k$.

 $\left(\hat{\lambda}_{j},\hat{\mathbf{b}}_{j};j=1,...,P\right)$ are eigenvalue-eigenvector pairs of $\mathbf{X}^{t}D\mathbf{X}$ this

leads to estimated principal components $\hat{Z}_j = \hat{\mathbf{b}}_j^t \mathbf{X}$

Estimators under consideration

 Linear Calibration estimator (Särndal & Lundström, 2005)

$$\hat{t}_{Lc} = \sum_{r} w_{k}^{pc} y_{k}$$

$$w_{k}^{pc} = d_{k} + d_{k} \boldsymbol{\delta}_{(pc)}^{t} \mathbf{Z}_{k} \quad \text{and} \quad \boldsymbol{\delta}_{(pc)} = \left(\mathbf{Z}^{t} \mathbf{D} \mathbf{Z}\right)^{-1} \left(\mathbf{T}_{z} - \mathbf{Z}^{t} \mathbf{d}\right)$$

$$\mathbf{D} = diag \left\{d_{1}, d_{2}, ..., d_{k}, ..., d_{m}\right\}$$

Propensity Score Calibration (Chang and Kott, 2008)

$$\hat{t}_{psCal} = \sum_{r} d_{k} \phi(\mathbf{x}_{k}^{t} \hat{\mathbf{\delta}}_{(pc)}) y_{k}$$
where:

$$\hat{\delta}$$
 is solution to $\sum_{r} d_{k} \phi(\mathbf{x}_{k}^{t} \hat{\delta}_{(pc)}) \mathbf{z}_{k} = \mathbf{T}_{z}$

Components Retention

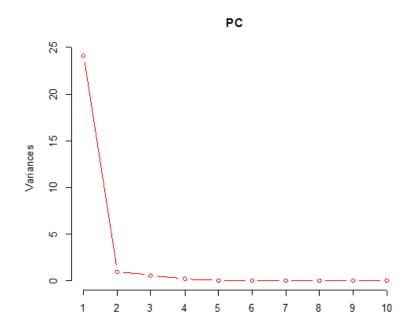
- Selecting R components among P is a theme that has widely been considered, e.g. Jolliffe (1972, 1973, 1982), Cadima and Jolliffe (1995), Jolliffe, Trendalov, and Uddin (2003), and McCabe(1984).
- A Canonical correlation between ${f H}$ and $\tilde{{f Z}}$

$$cor(\mathbf{H}, \tilde{\mathbf{Z}}) = \max_{\mathbf{P}_{\mathbf{H}}, \mathbf{P}_{\tilde{\mathbf{Z}}}} \frac{\mathbf{P}_{\mathbf{H}} \left(\mathbf{H}^{t} \tilde{\mathbf{Z}} \right) \mathbf{P}_{\tilde{\mathbf{Z}}}}{\left[\mathbf{P}_{\mathbf{H}} \left(\mathbf{H}^{t} \mathbf{H} \right) \mathbf{P}_{\mathbf{H}} \right]^{1/2} \left[\mathbf{P}_{\tilde{\mathbf{Z}}} \left(\tilde{\mathbf{Z}}^{t} \tilde{\mathbf{Z}} \right) \mathbf{P}_{\tilde{\mathbf{Z}}} \right]^{1/2}}$$

- H vector of model variables
- $\tilde{\mathbf{Z}}$ the portion of \mathbf{Z} in to response set

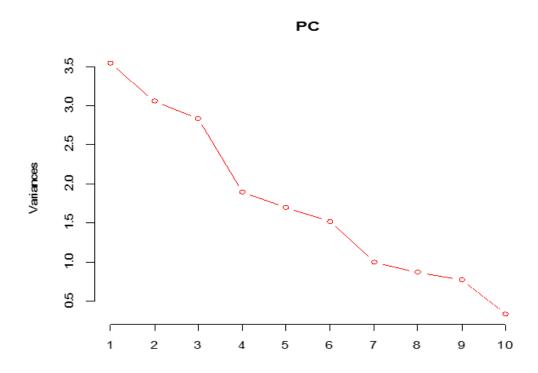
Simulation

- The use of PCs is illustrated in two simulation studies:
- Study 1 with auxiliary variables having strong correlation structure as the scree plot below illustrates



Simulation-cont.

 Study 2 where auxiliary variables are weakly correlated as the scree plot illustrates



Results

Table 1: LC on original population auxiliary variables vs. LC on population PCs –Study1

	Properties	Estimators		
Sample size		L. Calibration on X	L. Calibration on PCs	
	Rel.bias(%)	5.474	1.296	
300	S.E.	3519	935	
	RMSE	8661	2094	
	Rel.Bias(%)	3.974	1.149	
600	S.E.	3135	846	
	RMSE	6544	1864	

Table 2: LC on original Sample auxiliary variables vs. LC on Sample PCs –Study1

		Estimators			
Sample size	Properties	L. Calibration on X	L. Calibration on PCs		
	Rel.bias(%)	3.930	0.192		
300	S.E.	21,936	11,202		
	RMSE	22,660	11,206		
	Rel.Bias(%)	2.951	0.369		
600	S.E.	12,422	7608		
	RMSE	13,134	7626		

Table 3: LC on original population auxiliary variables vs. LC on population PCs-Study 2

		Estimators		
Sample size	Properties	L. Calibration on X	L. Calibration on PCs	
	Rel.bias(%)	0.033	0.010	
300	S.E.	5767	5588	
	RMSE	5671	5588	
600	Rel.Bias(%)	0.009	0.025	
	S.E.	3769	3947	
	RMSE	3769	3950	

Table 4: LC on original Sample auxiliary variables vs. LC on Sample PCs-Study 2

		Estimators		
Sample size	Properties	L. Calibration on X	L. Calibration on PCs	
	Rel.bias(%)	0.007	0.006	
300	S.E.	5733	5645	
	RMSE	5733	5645	
	Rel.Bias(%)	0.024	0.026	
600	S.E.	3913	3996	
	RMSE	3914	4000	

Table 5: PS on original population auxiliary variables vs PS on population PCs-Study 1

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		Estimators				
Sample size	Properties	PS on X	Time (in hr)	PS on PCs	Time (in hr)	
	Rel.bias(%)	0.280		0.153		
300	S.E.	16,182	7	15,912	0.35	
	RMSE	16,188		15,914		
	Rel.Bias(%)	0.169		0.264		
600	S.E.	10,899	36	10,757	1.30	
	RMSE	10,902		10,764		

Table 6: PS on original Sample auxiliary variables vs PS on Sample PCs-Study 1

		Estimators				
Sample size	Properties	PS on X	Time (in hr)	PS on PCs	Time (in hr)	
	Rel.bias(%)	0.255		0.125		
300	S.E.	16,161	0.25	16,010	0.18	
	RMSE	16,166		16,011		
	Rel.Bias(%)	0.191		0.263		
600	S.E.	10,880	0.50	10,795	0.25	
	RMSE	10,884		10,801		

Table 7: PS on original population auxiliary variables vs. PS on population PCs-Study 2

		Estimators		
Sample size	Properties	PS on X	PS on PCs	
	Rel.bias(%)	0.587	0.670	
300	S.E.	18,522	19,738	
	RMSE	18,894	20,212	
600	Rel.Bias(%)	0.106	0.140	
	S.E.	6035	6781	
	RMSE	6072	6839	

Table 8: PS on original Sample auxiliary variables vs. PS on Sample PCs-Study 2

		Estimators		
Sample size	Properties	PS on X	PS on PCs	
	Rel.bias(%)	0.196	0.264	
300	S.E.	10452	12222	
	RMSE	10526	12334	
600	Rel.Bias(%)	0.002	0.006	
	S.E.	4155	4167	
	RMSE	4155	4167	

Table 9: Estimated model coefficients (Population auxiliary information-Study 2)

 δ_0 δ_1 δ_2

True Coefficients

1.306, -0.020, -0.083

Sample size		PS on X			PS on PCs	
	$\hat{\delta}_{_{0}}$	$\hat{\delta}_{_{1}}$	$\hat{\delta}_{_{2}}$	$\hat{\delta}_{_{0}}$	$\hat{\delta}_{_{1}}$	$\hat{\delta}_{_{2}}$
300	1.128 (1.478)	-0.017 (0.000)	-0.036 (0.062)	1.182 (1.413)	-0.019 (0.000)	-0.039 (0.059)
600	1.205 (0.578)	-0.018 (0.000)	-0.066 (0.024)	1.238 (0.656)	-0.018 (0.000)	-0.069 (0.030)

Conclusion

- The results suggest the use of PC to be effective as this does not distort the results
- PCs are effective than original auxiliary variables in the conditions of the study in terms of computational effort.
- When the correlation structure is strong PCs are effective in PS calibration scheme than in LC scheme while weak correlation structure in auxiliary variables turns PCs more effective in LC than in PS

Thank you very much for your attention!